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Measuring Total Factor
Productivity at the Firm
Level using OECD-ORBIS

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ABSTRACT/RÉSUMÉ

Measuring Total Factor Productivity at the Firm Level using OECD-ORBIS

Recent OECD research has utilised harmonised cross-country firm level data to explore the contribution of public policies to cross-country differences in productivity, innovation and resource allocation. This paper describes the steps taken to and the trade-offs involved in constructing firm-level total factor productivity (TFP) measures using ORBIS, a cross-country longitudinal firm-level database available from Bureau van Dijk, an electronic publishing firm. First, it shows that not all productivity measures can be calculated using readily available variables for all countries, and presents possible solutions to this problem by using imputations for certain variables. Second, it assesses the accuracy of these imputations on a set of countries where the available data in ORBIS provides a good coverage, for a wide range of TFP measures. Indeed, an extensive comparison of the actual and the imputed values of TFP for those countries suggests that TFP measures using imputations provide a reasonable approximation for the "true" values. Furthermore, to improve representativeness, resampling weights are constructed - which help correcting for the underrepresentation of small firms - while for the sake of international comparability, industry-level PPP conversions are also applied. Finally, as a plausibility check and to illustrate the potential of the database, the paper explores the country-composition of the globally most productive firms, the forces of convergence to the productivity frontier and the impact of regulation on productivity growth, in a sample of 18 OECD countries.

JEL classification codes: D24, O47, D22

Keywords: Productivity measurement; firm-level data; cross-country analysis

Mesurer la productivité totale des facteurs au niveau de l'entreprise à l'aide de la base de données OCDE-ORBIS

Des travaux récents de l'OCDE ont utilisé des données harmonisées d'entreprises des pays de l'OCDE afin d'étudier la contribution des politiques publiques aux différences entre pays dans la productivité, l'innovation et l'allocation des ressources. Ce document décrit les mesures prises pour et les compromis impliqués dans la construction des séries au niveau de l'entreprise, de la productivité totale des facteurs (PTF) à l'aide d'ORBIS, une base de données d'entreprises longitudinale mises à disposition par le Bureau van Dijk, une maison d'édition électronique. D'abord, il montre que toutes les mesures de la productivité ne peuvent être calculées à l'aide de variables facilement disponibles dans tous les pays, et présente d'éventuelles solutions à ce problème en utilisant des imputations pour certaines variables. Deuxièmement, il évalue l'exactitude de ces imputations sur un ensemble de pays où les données, disponibles dans ORBIS, offrent une bonne couverture pour un large éventail de mesures de la PTF. En effet, une comparaison approfondie des valeurs réelles et imputées des PTF pour ces pays suggère que les mesures de la PTF, utilisant des imputations, constituent une approximation raisonnable des valeurs «réelles». En outre, afin d'améliorer la représentativité, des poids de rééchantillonnage ont été construits – afin d'aider à corriger la sous-représentativité des petites entreprises - pour des raisons de comparabilité internationale, des conversions en PPP ont également été appliquées au niveau des industries. Enfin, comme contrôle de plausibilité et pour illustrer le potentiel de la base de données, ce document examine la composition pays des entreprises les plus productives au niveau mondial, les forces de convergence vers la frontière de la productivité et l'impact de la réglementation sur la croissance de la productivité, dans un échantillon de 18 pays de l'OCDE.

Codes JEL : D24, O47, D22

Mots clés : mesure de la productivité ; données sur les entreprises ; analyses comparatives entre pays

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Measuring Total Factor Productivity at the Firm Level using OECD-ORBIS

Peter N. Gal*

17th May 2013

1 Introduction

Important questions regarding the determinants of firm-level productivity, reallocation and firm dynamics can only be answered if one uses firm-level data.¹ Similarly, analyses based on firm-level data have the potential to more credibly identify the effects of certain policies than studies using only aggregate (country or industry-level) data, and also to describe the mechanisms behind the policy effect in more detail. Accordingly, accessing high-quality data at the firm-level is a priority for many researchers in the field of firm behavior and productivity. While the most comprehensive firm-level data sources are usually held by national statistical agencies or tax offices (e.g. business registers, production surveys, tax returns), data access restrictions due to confidentiality concerns are key barriers in cross-country research.² Reflecting these constraints, researchers often turn to commercial data sets, such as ORBIS, which contain more timely cross-country firm-level information and are more easily accessible than official, statistical sources.³

Recent OECD research has utilized harmonized cross-country firm level data from ORBIS to explore the contribution of public policies to cross-country differences in productivity, innovation and resource allocation (see Andrews and Cingano, 2012; Andrews et al., 2013; Andrews and Criscuolo, 2013). Before this research was undertaken, however, considerable effort was put into improving the reliability and international comparability of these firm level data, and to address a number of complicated methodological issues – outlined below – that are posed by the design of the database. Against this backdrop, this paper describes these efforts as well as the specific trade-offs involved when using ORBIS for productivity

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¹For surveys on heterogeneity of productivity across firms and within sectors, see Bartelsman and Doms (2000) and Syverson (2011). For the aggregate productivity effects of the allocation of resources, see Hsieh and Klenow (2009) and Bartelsman et al. (2009). For a related literature on adjustment costs, see Hamermesh and Pfann (1996).

²A possible workaround to this problem is *distributed microdata analysis*, developed by Eric Bartelsman (see Bartelsman et al. (2005, 2009), among others) and has been also used by the studies in the OECD (Bartelsman et al., 2003 and more recently Biosca et al. (2012)). It involves a decentralized analysis by collecting micro-aggregated data (e.g. by industry, firm size, age, etc.) on a set of economic variables from official, national firm-level data sources, which are produced by standard program codes sent to local experts. Despite its main advantage of using the widest possible coverage and best data quality, its drawbacks, such as the inability to pool together the data set and run cross-country firm-level regressions in one site, the time-consuming nature of the procedure and publication lags, may make it a less attractive choice in some cases than using commercial data such as ORBIS.

³The ORBIS version used in this report is the one downloaded by the OECD from the data vendor Bureau van Dijk in June 2011. For more details, see Section 2 and Gonnard and Ragoussis (2013).

analysis. It builds on earlier work done by the OECD on the ownership structure and on cross-country harmonization (see Gonnard and Ragoussis, 2013), with a more specific focus on productivity measurement and on further improvements of international comparability (e.g. resampling and PPP-conversion). Moreover, it carries out a wide range of quality and robustness checks of the measures obtained, also highlighting various potential uses of the data. As such, it aims to serve as a reference for current and future firm-level analyses utilizing ORBIS either within or outside the OECD.

ORBIS brings together firm-level data from many countries, thus enabling firm-level cross-country analysis. Research using Amadeus, the earlier, “European-version” of ORBIS, has explored the effects of entry regulations on business dynamics (Klapper et al., 2006), product market regulations on productivity growth (Arnold et al., 2008) and corporate taxes on investment (Schwellnus and Arnold, 2008). In addition to the European countries included in Amadeus, ORBIS covers other, non-European countries, notably the United States, Japan and Korea, and some important emerging countries like Brazil, China and Russia. Despite this potential to extend firm-level analysis into a global context, ORBIS does not cover all variables equally well in all countries. Hence there is a trade-off between coverage (i.e. the set of firms and countries included in the analysis) and the type of analysis one can carry out. For instance, there are some limitations regarding data from the United States in ORBIS, but it is important – from the perspective of productivity analysis – that these issues are addressed, since firms from the United States are likely to be at the technological frontier in many industries.

More specifically, the key obstacle to including a wider coverage of countries in the analysis is the lack of availability of certain variables (value added and intermediate inputs) required for measuring total factor productivity (TFP). Hence, this paper – continuing along the lines of Gonnard and Ragoussis (2013) – examines the scope for imputing these variables, using both available information in ORBIS and from industry-level data sources (such as the Structural Analysis Database of the OECD, or STAN). To assess the appropriateness of imputations, the paper compares TFP measures using imputed and non-imputed variables for those countries with the most comprehensive data coverage.

Across a variety of TFP measures, reasonably high correlations, in the range of 0.7-0.95, are observed between the measures based on actual values and on imputed values, for levels as well as growth rates. Thus, for instances where imputations are required (e.g. the United States), TFPs based on imputed values may serve as a reasonable approximation for the true unobserved values. This results in an increase in the cross country sample by about 60%.⁴

In order to correct for the underrepresentation of small firms, resampling weights are also constructed using the total number of employees for each country, industry, year and size-class from the OECD Structural and Demographic Business Statistics (SDBS). Moreover, for the sake of international comparability, industry-level PPP conversions are applied.

A final database of 24 countries results, once these procedures have been implemented. As a plausibility check and to illustrate the potential uses of the final database, the paper explores which countries are most successful in reaching the global productivity frontier, the forces of convergence to the frontier and the impact of product market regulation in upstream sectors on productivity growth.

Section 2 gives a detailed outline and defines key concepts, and explores (i) data coverage issues; (ii) the potential use of imputations to enhance coverage; (iii) the construction of resampling weights and (iv) price indices. Section 3 discusses measurement and estimation of TFP at the firm-level and a set of general practical issues for users of the data. Section 4 assesses the reliability of TFP measures based on imputed data. Using the constructed data set, Section 5 presents a few cross-country applications, and Section 6 offers some conclusive thoughts.

⁴For some countries (Japan and Korea, Norway, Switzerland and the United States), there are some qualifications, see Table 5.

2 Preparing ORBIS for TFP Measurement

2.1 Detailed Outline and Key Concepts

A harmonized cross-country database on firm-level TFP, inputs and output is a useful tool for analysing a range of policy questions related to productivity, and since TFP controls for both capital and labor (and possibly intermediate inputs), it has the potential to more accurately describe the efficiency with which resource are used than labor productivity. However, there is no single best measure for TFP, some variants may be better suited for certain purposes than others (see Section 3 and the Box therein on productivity measurement). Bearing in mind these differences, the paper takes four approaches to calculate TFP: the simplest Solow residual (least demanding in terms of data availability and quality), the superlative index (best suited for level comparisons), OLS-residuals (benchmark for production function based methods) and the recent Wooldridge (2009) method (most sophisticated but most demanding in terms of data).

Numerous steps are initially taken in order to enhance the international comparability of the measurement. First, to increase the country coverage as well as the number of firms within countries, imputation methods are introduced for certain variables (value added and intermediate inputs). Next, in order to correct for the typical under-representation of small firms, resampling weights are constructed using administrative sources. As final steps in the data preparation, to ensure comparability of productivity levels across countries and over time, currency conversion based on PPPs and deflation are discussed, and capital stocks are also constructed. The imputation methods and the resampling weights build on and refine earlier work by Gonnard and Ragoussis (2013).

In the next step, the actual estimation of TFP is carried out, for the 4 types of measures mentioned above, and labor productivity is also calculated as a benchmark. The details of these various measures and a general background on productivity measurement is given in Section 3, while the characteristics of the resulting database (country and industry coverage) as well as its limitations are also discussed.⁵ In turn, in order to test the quality of the imputations, detailed comparisons for productivity levels, growth rates and their within-industry distributions are carried out for the set of observations where non-imputed and imputed estimates are both available, for all the productivity measures considered (Section 4). To give a sense of the similarities and differences between the various productivity measurement approaches, they are also compared using non-imputed data.

Section 5 highlights some illustrations for the uses of the data, exploiting the wide cross-country coverage (up to 24 countries) and the enhanced international comparability both in terms of a more balanced coverage (by using resampling weights) and in productivity levels (by using industry level PPPs).

2.2 Coverage of ORBIS

For some countries in ORBIS, firm-level TFP measures can only be calculated if one is prepared to use additional information, i.e. *imputations*, for value added and in some cases, intermediate inputs (see Section 2.3 below for details on imputation methods). To illustrate the problem, Table 1 presents the number of available observations per country where labour productivity (using turnover as a measure of output) and the TFP measures can be constructed, for a selected group of countries.⁶ The table ranks countries according to the ratio of the number of observations where TFP can be measured and

⁵The data behind all the tables and calculations below concentrates on the market sector (NACE Rev 1.1 codes 15-74, i.e. all industries except agriculture, mining and public services) for the years 1999-2009, if not stated otherwise, and is derived from filtering out accounts referring to less than complete calendar years and consolidated accounts, using indicator variables prepared Gonnard and Ragoussis (2013).

⁶See Table E.1 in the Appendix for the same table with all countries and with more sophisticated TFP measures, and also for 2002 in Table E.2.

Table 1: Comparing the availability of variables for productivity measures for selected countries
Number of observations by country and by productivity measures, for the years 1999-2009

Country	Labor Prod. using Turnover	TFP, non imputed	Non-imputed TFP / Labor productivity	TFP, internally imputed	Internally imputed TFP / Labor productivity	TFP, externally imputed*	Externally imputed TFP / Labor productivity*
ITA	2,010,555	1,748,028	86.9%	1,843,820	91.7%	1,944,699	96.7%
ESP	4,383,651	3,675,628	83.8%	3,945,899	90.0%	4,222,513	96.3%
GBR	542,875	307,595	56.7%	520,661	95.9%	555,392	102.3%
FRA	4,034,693	1,556,908	38.6%	2,877,769	71.3%	3,134,441	77.7%
JPN	1,432,137	460,628	32.2%	515,308	36.0%	867,051	60.5%
DEU	917,426	171,408	18.7%	170,734	18.6%	174,980	19.1%
EST	210,969	129	0.1%	140,130	66.4%	153,124	72.6%
GRC	184,634	12	0.0%	13	0.0%	162,909	88.2%
USA	19,392,887	8	0.0%	16	0.0%	29,142	0.2%
Total (OECD)	42,282,214	10,840,462	25.6%	15,167,998	35.9%	17,027,212	40.3%
Mean (OECD)	1,243,595	318,837	25.1%	446,118	59.2%	500,800	75.1%

Note: Countries are ranked, in a decreasing order, by the coverage of the sample for which the simplest TFP-measure (e.g. Solow residual) without imputations can be calculated. The selected country groups are intended to represent a broader set of countries with either good coverage (top three countries), poor coverage (bottom three countries) or medium coverage (group in the middle).

The columns referring to TFP show the number of observations where TFP can be calculated, using the following variables (ORBIS variables in parentheses): value added (ADDED_VALUE), capital (TANGIBLE_FIXED_ASSETS and DEPRECIATION), employment (EMPLOYEES). For details on the exact definition of internally imputed TFP-s (using only ORBIS) and externally imputed ones (using also industry-level, external sources), see Table 2.

The data underlying the table refers to the non-farm business sector (NACE Rev 1.1 codes 15-74) for the years 1999-2009, and is derived from filtering out accounts referring to less than complete calendar years and consolidated accounts, using indicator variables prepared by the Statistics Directorate (STD) of the OECD (Gonnard and Ragoussis, 2013). The actual numbers in the TFP database may be lower than here if price indices are not available or other data problems prevent the calculation of TFP measures for some firms. The United States, although giving close to 30 thousand potentially useful observations for productivity calculations, is still limited due to the low number of observations with profits (earnings before interest, taxes, depreciation and amortization, EBITDA) and capital. See more on this in Section 3.1.1.

* The numbers in the column *External imputations* assume administrative, industry level data on labour costs and employment is available in each industry and country.

the number of observations where turnover-based labour productivity can be calculated. Generally, most large European countries, such as Spain, Italy, and Great Britain, but also smaller ones like Sweden and Portugal, have data which are almost equally well suited for obtaining measures of either labour productivity or TFP. By contrast, for some English-speaking countries (United States, Canada) and small EU countries (Greece, Estonia), the variables required for TFP calculations are not readily available, or only in very small numbers.

For a number of countries, data coverage can be extended via the use of data imputations. The use of internal imputations, which exploits information on other firm-level variables in ORBIS, can significantly increase the number of firms included in the TFP sample for Estonia, France and Great Britain. However, for countries such as the United States and Greece, external imputation methods, i.e. those which rely in part on external, industry level data from OECD STAN, are required to yield an adequate number of observations (see Table 2 for details). The use of external imputation methods also increases the available number of firms for TFP measurement for a number of other countries (e.g. Denmark, Estonia, Japan, the Netherlands, Norway and Portugal, as shown in Table E.1 of the Appendix).

The last row of Table 1 shows that for the *average* OECD country in ORBIS, the coverage of the non-imputed TFP sample to the labour productivity sample is 25.1%. The coverage increases to 59.2% and to 75.1% when using internal and external imputations, respectively. For the total number of observations in all OECD countries pooled (the row *Total*), the sample goes up from 25.6% to 35.9% and 40.3%,

respectively.⁷

2.3 Extending the Coverage: Using Imputations

Table 2 contains information on the specific data requirements for productivity calculations and highlights which values can be imputed in order to extend the coverage. By far the most important variable where imputations help is value added. In some countries, it can be imputed *internally*, that is, using other variables in ORBIS. Based on correlations across levels and growth rates (see Gonnard and Ragoussis (2013)), the best substitute for value added (VA) is simply using its definition based on factor incomes. This entails adding up factor incomes going to employees (total wage bill, wL , i.e. average labour w cost times the number of employees L) and to capital owners (profits, rK , i.e. the rate of return r times the capital stock K):

$$VA = wL + rK \quad (2.1)$$

The empirical counterparts to these variables in ORBIS are the COSTS_EMPLOYEES and EBITDA (Earnings Before Interest Taxes Depreciation and Amortisation), where the latter is a measure of profits (i.e. part of the income going to capital):⁸

$$VA_{csit}^{\text{InternallyImputed}} = COSTS_EMPLOYEES_{csit}^{ORBIS} + EBITDA_{csit}^{ORBIS}, \quad (2.2)$$

where subscripts c, s, i, t denote country, sector, firm and time, respectively, and superscripts denote the source of the variable. Such internally imputed values are essentially alternative definitions of firm value added, and indeed show very high correlations for the log-levels (0.98 for an average country and year) and growth rates (around 0.8).⁹

However, in some countries (e.g. the United States), such internal imputations cannot be obtained because of missing labour cost data.¹⁰ Thus imputations based on external data are required (*external imputation*). Following Gonnard and Ragoussis (2013), labour costs are imputed by relying on industry level data from OECD STAN. First, by dividing the total labour cost by the number of employees per country, year and 2-digit industry using STAN aggregates, average labour cost per worker is obtained, \bar{w}_{cst}^{STAN} . This is then multiplied by the number of employees in ORBIS (firm-by-firm and year-by-year) in order to get imputed firm- and year specific labour costs. Finally, this imputed labour cost is added to observed EBITDA-s, resulting in value added measures which vary over firms and time:

$$VA_{csit}^{\text{ExternallyImputed}} = \bar{w}_{cst}^{STAN} * L_{csit}^{ORBIS} + EBITDA_{csit}^{ORBIS}. \quad (2.3)$$

This approach is expected to work well if within-industry wage differentials are not too high.¹¹ However, given the existence of a firm-size premium on wages (Brown and Medoff, 1989; Oi and Idson, 1999), primarily driven by large firms' higher productivity (Idson and Oi, 1999), it is important to address this potential problem.¹² In order to capture part of the dispersion in wages, wage differentials from the

⁷Also, for some countries, TFP measures using imputations can actually bring more observations than using turnover based labour productivity, hence some ratios are larger than 100%. Note also that these numbers are based on increases in the availability of variables needed for TFP calculations, but the availability of price indices or other data problems may constrain it. See Table 5 for more details on this issue.

⁸The other possibility for internally imputing value added – taking the difference between output and intermediate inputs (row *Internally imputed value added (2)* in Table 2) – does not help in practice, because intermediate inputs are only partially available in ORBIS. An important part of it, material costs are captured, but that is also not widely available.

⁹See the entries referring to internal imputations, shown in Table E.3 in the Appendix.

¹⁰It is available only in about 10 cases, as mentioned also in Gonnard and Ragoussis (2013).

¹¹A priori, however, it is difficult to say what is "too high" from the point of view of productivity estimates, hence it is an empirical issue.

¹²The early estimates for the United States concluded that "an employee working at a location with $\ln(\text{employment})$ one standard deviation (which equals about two) above average can be expected to earn 6-15 percent more than a similar

industry average are predicted using a quadratic function of firm size, profit per employee and age as explanatory variables, for each country and two digit industry.¹³ Such predictions, \hat{w}_{csit}^{ORBIS} , serve as the basis for imputing the firm-level wages in the second type, *refined external imputations*. Formally, they are constructed as follows:

$$VA_{csit}^{\text{ExternallyImputed2}} = \hat{w}_{csit}^{STAN,ORBIS} * L_{csit}^{ORBIS} + EBITDA_{csit}^{ORBIS}, \quad (2.4)$$

where $\hat{w}_{csit}^{STAN,ORBIS}$ is the sum of the mean industry value of wages \bar{w}_{cst}^{STAN} (taken from OECD STAN) and the predicted wage-differentials from the mean \hat{w}_{csit}^{ORBIS} :

$$\hat{w}_{csit}^{STAN,ORBIS} = \bar{w}_{cst}^{STAN} + \hat{w}_{csit}^{ORBIS}. \quad (2.5)$$

Correlations between the non-imputed and the externally imputed value added measures confirm that the approximation works well: correlations are higher than 0.9 for log-levels and around 0.6 for growth rates (see Table E.3 of the Appendix) and they improve when within-industry wage-differentials are taken into account. In Section 4.1, such correlations are calculated directly for the TFP measures, among other checks, in order to explore if TFP measures based on these value added imputations serve as a reasonable approximation for those measures which are based on non-imputed data.¹⁴

2.4 Improving Representativeness

Smaller and younger firms, especially in services sectors, are typically underrepresented in ORBIS.¹⁵ Therefore, in order to improve representativeness along the sectoral and firm size dimensions, sampling weights are introduced, using information on the number of employees in a country*industry*size-class*year cell from the OECD SDBS database, which is based on official sources and the full population of businesses. In short, a time-varying resampling weight is assigned to each firm, which is always greater or equal to one, making sure we do not lose firms which are already in ORBIS, but we only *replicate* them up to the point where the true size- and sectoral structure is achieved.^{16,17} The procedure is described in detail below.

employee at a location with ln(employment) one standard deviation below average.” (Brown and Medoff, 1989). I am grateful to Chiara Criscuolo for calling attention to this issue.

¹³Profit per employee is a proxy in these regressions for labour productivity, as labour productivity is not always available when profits are. The age of the firm is constructed by subtracting the incorporation year from the actual year. The R^2 -s from these predictions are in the range of 0.3-0.5, depending on the country (detailed results are available upon request). For countries where no labour cost data is available, a country with similar labour market institutions is chosen to estimate the coefficients which serve as the basis for predictions. For the United States, Great Britain is chosen, for Greece, two countries, Spain and Portugal are chosen, based on the proximity of employment protection legislation indicators described in Venn (2009).

¹⁴Of course, in any analysis, each firm and country should use the same type of imputation (only internal, only external or no imputation), i.e. it is not advisable to mix imputed and non-imputed TFP measures in a single regression, for instance. Accordingly, in the resulting ORBIS-TFP database, several TFP variables are constructed, each of them calculated in a uniform way: either using only non-imputed data or only the same type of imputation, for each firm-year observation (see more on this in the description of the final productivity database in Section F of the Appendix).

¹⁵See Table 3 below. For countries with low number of observations, this problem is relatively more important, and this issue can also hamper international comparability.

¹⁶In software packages such as Stata, some calculations (e.g. standard deviations, correlations, within-group percentiles and panel regressions) cannot be carried by using non-integer weights, but only integer ones - either as an option (e.g. *fweight* in Stata) or in a two-step manner, after expanding the database with respect to the weight variable.

¹⁷Experiments were also performed using the number of firms or establishments instead of the number of employees as the the match variable at each country*industry*size-class*year cell. The number of employees was chosen, because it avoids problems of some countries reporting the number of establishments (Japan, Korea) while the rest the number of firms (enterprises), and in ORBIS, the unit of observation is usually the firm (enterprise). An exception is the United States, where it is the plant (establishment) which is the unit of observation. This is taken into account by using only single-establishment firms from the United States database, as multi-establishment firms currently do not have revenue or sales information at this point in the OECD-ORBIS database. The Statistics Directorate (STD) of the OECD is working on a solution and in future versions of the database, large multinational firms from the United States may be included as well. Despite this issue, many large and highly productive firms are observed for the United States, see Tables 3 and 10.

Table 2: Imputations and variable definitions

Variable	Imputation	ORBIS variable(s) involved (short name)	External variable(s) involved*
Gross output, or total revenue, TR	-	$OPERATING_REV_TURNOVER^{**}$ (TR)	Output price index
Value added, VA	-	$ADDED_VALUE$ (VA)	Value added (VA) price index
Internally imputed value added , VA^{int}	$VA^{int} = wL + rK$	$COSTS_EMPLOYEES$ (wL), $EBITDA$ (rK)	VA price index
Internally imputed value added (2), $VA^{int,2***}$	$VA^{int,2} = TR - II$	$OPERATING_REV_TURNOVER$ (TR), $MATERIAL_COSTS$ (II)	VA and interm. input price index
Externally imputed value added , VA^{ext}	$VA^{ext} = w^{STAN}L_{ORBIS} + rK$	$EMPLOYEES$ (L), $EBITDA$ (rK)	$\frac{LabourCosts}{Employment} +$
Externally imputed value added (2) , $VA^{ext,2}$	$VA^{ext,2} = \hat{w}^{STAN}L_{ORBIS} + rK$	$EMPLOYEES$ (L), $EBITDA$ (rK), $EMPLOYEES$ (L)	$\frac{LabourCosts}{Employment} +$
Employment, L	-	$EMPLOYEES$ (L)	-
Capital, K ++	-	$TANGIBLE_FIXED_ASSETS$ (K), $DEPRECIATION$ ($DEPR$)	Investment price index
Intermediate inputs, II +++	-	$MATERIAL_COSTS$ (II)	Intermediate impute price index
Imputed intermediate inputs, II^{int}	$II^{int} = TR - VA$	$OPERATING_REV_TURNOVER$ (TR), $ADDED_VALUE$ (VA)	Intermediate impute price index
labour cost share, s_L	$s_L = \frac{wL}{VA}$	$COSTS_EMPLOYEES$ (wL), $ADDED_VALUE$ (VA)	-
Externally imputed labour cost share, s_L^{ext}	$s_L^{ext} = \frac{(wL)^{STAN}}{(VA)^{STAN}}$	-	labour costs and nominal value added

* Price indices are always from external sources (OECD STAN for changes over time and *Timmer et al. (2007)* or average 2005 euro/local currency exchange rates for comparisons across countries).

See details in Section B of the Appendix.

** The variable available in ORBIS is not exactly gross output, but it is usually substituted for revenues or sales in firm-level databases (See, for instance, Foster et al., 2001). The ORBIS variable *SALBS* is only scarcely available and when available, it is nearly identical to *OPERATING_REV_TURNOVER*, hence this latter variable is used as a measure for gross output.

*** Unfortunately, this definition of value added is not useful for imputations in the case of ORBIS, because where value added is missing, usually intermediate inputs are missing as well. Also, note that only part of intermediate inputs are captured by the ORBIS variable *MATERIAL_COSTS*.

+ When using external imputations for the cost of employees, total labour costs / employees are taken from OECD STAN (per country, year, 2-digit industry) and multiply it with the firm-level, time-varying number or employees. For the second type of externally imputed value added, see the definition of \hat{w} in the main text.

++ For calculating the capital stock by using the standard Perpetual Inventory Method (PIM), see Section A of the Appendix.

+++ As no energy usage and purchased services are available in ORBIS, material costs (*ORBIS* variable: *MATERIAL_COSTS*) are used as a proxy for total intermediate inputs.

Note: superscript *int* and *ext* mean internally (using ORBIS) and externally (using STAN) imputed values, respectively.

First, the latest available data from OECD.STAT is used on the number of employees, by country, industry, firm size class and year. Since the database is not balanced, i.e. not all dimensions are available for each country and year, some reasonable assumptions are needed to “fill up” the missing cells. For details, see Section C of the Appendix.

With the resulting SDBS database, the total number of employees, L_{cist}^{SDBS} , is compared to what is obtained from ORBIS, L_{cist}^{ORBIS} , in the same country * industry * sizeclass * year cell (denoted by subscripts c, i, s , and t , respectively). In turn, a weight variable w_{jt} is constructed for each firm j and year t , such that the expected value of a weight will be determined by the ratio of SDBS to ORBIS employment within the cell:

$$E(w_{jt}) = \frac{L_{cist}^{SDBS}}{L_{cist}^{ORBIS}}, \quad (2.6)$$

where L denotes the number of employees, and the expected value $E(\cdot)$ is taken over those firms which belong to the same c, i, s cell, for each year t . The logic behind such a weighting variable is that it “scales up” ORBIS observations in a cell so that they match those observed in SDBS.¹⁸ Note that depending on the set of variables one focuses on, L_{cist}^{ORBIS} may also differ, since the availability of those in ORBIS may differ. As a result, the resampling weights will differ, and thus in the constructed database, a corresponding resampling weight is calculated for each productivity measure.¹⁹

The amount of resampling needed per country and firm size class is shown in Table 3. As expected, the smallest firm-size classes tend to be more under-represented, whereas the larger ones usually do not need weights much larger than 1. Also, concentrating on the set of firms with more than 20 employees helps in getting a better coverage. In some countries (e.g. Belgium, Estonia, Finland, and Sweden), resampling weights are very close to one, across all firm sizes, indicating that the structure and coverage of ORBIS is close to that of the true population of firms. Regarding the industry dimension (shown in Table E.5 of the Appendix), the services sectors have on average 2-3 times higher resampling weights than manufacturing. Energy, water and gas, and also construction are the most well covered and need the least amount of scaling up.

Figure 2.1 shows the improvement of ORBIS in matching aggregate employment figures when using the constructed resampling-weights compared to the case when only unweighted employment is used. Furthermore, Figure 2.2 shows the average number of employees when the resampling weights are applied. The numbers look sensible given that large developed economies tend to have larger firms (Great Britain, United States), and Mediterranean European countries (Italy, Portugal, Spain) tend to have smaller firms.

¹⁸It is implemented by assigning a weight to each firm-year observation which is always greater or equal to one, making sure we do not lose firms which are already in ORBIS. For instance, if the ratio $\frac{L_{cist}^{SDBS}}{L_{cist}^{ORBIS}}$ is 1.3, that is, SDBS employment is 30% higher than ORBIS employment, then the 30% “extra” employment is obtained by drawing firms randomly from the pool of ORBIS firms, such that the “extra” firms will make up for the missing 30%. In a different case, if the ratio is $\frac{L_{cist}^{SDBS}}{L_{cist}^{ORBIS}}$ is 2.3, then all firms are taken at least twice and only the remaining extra 30% will be drawn randomly. In this way, it is ensured that the resampled pool is as close to the original ORBIS pool as possible, and the random selection plays only a limited role in changing the composition of firms within a cell.

Formally, the actual resampling weights \tilde{w}_{jt} are obtained by

$$\tilde{w}_{jt} = 1 + \left\lfloor \frac{L_{cist}^{SDBS} - L_{cist}^{ORBIS}}{L_{cist}^{ORBIS}} \right\rfloor + z_{jt}$$

where z_{jt} is a Bernoulli distribution (having values of either zero or one: $z_{jt} \in \{0, 1\}$) with expected value $Pr(z_{jt}) = \frac{\text{mod}(L_{cist}^{SDBS}, L_{cist}^{ORBIS})}{L_{cist}^{ORBIS}}$ (i.e. the remainder of $\frac{L_{cist}^{SDBS}}{L_{cist}^{ORBIS}}$, which is, in the example above, 30%), and $\lfloor x \rfloor$ denotes the floor of x (i.e. the largest integer which is smaller or equal to x). This construction will ensure that the actual values of \tilde{w}_{jt} will be integers and always larger or equal to one, as well as their sample mean in a given cell will be close to $\frac{L_{cist}^{SDBS}}{L_{cist}^{ORBIS}}$.

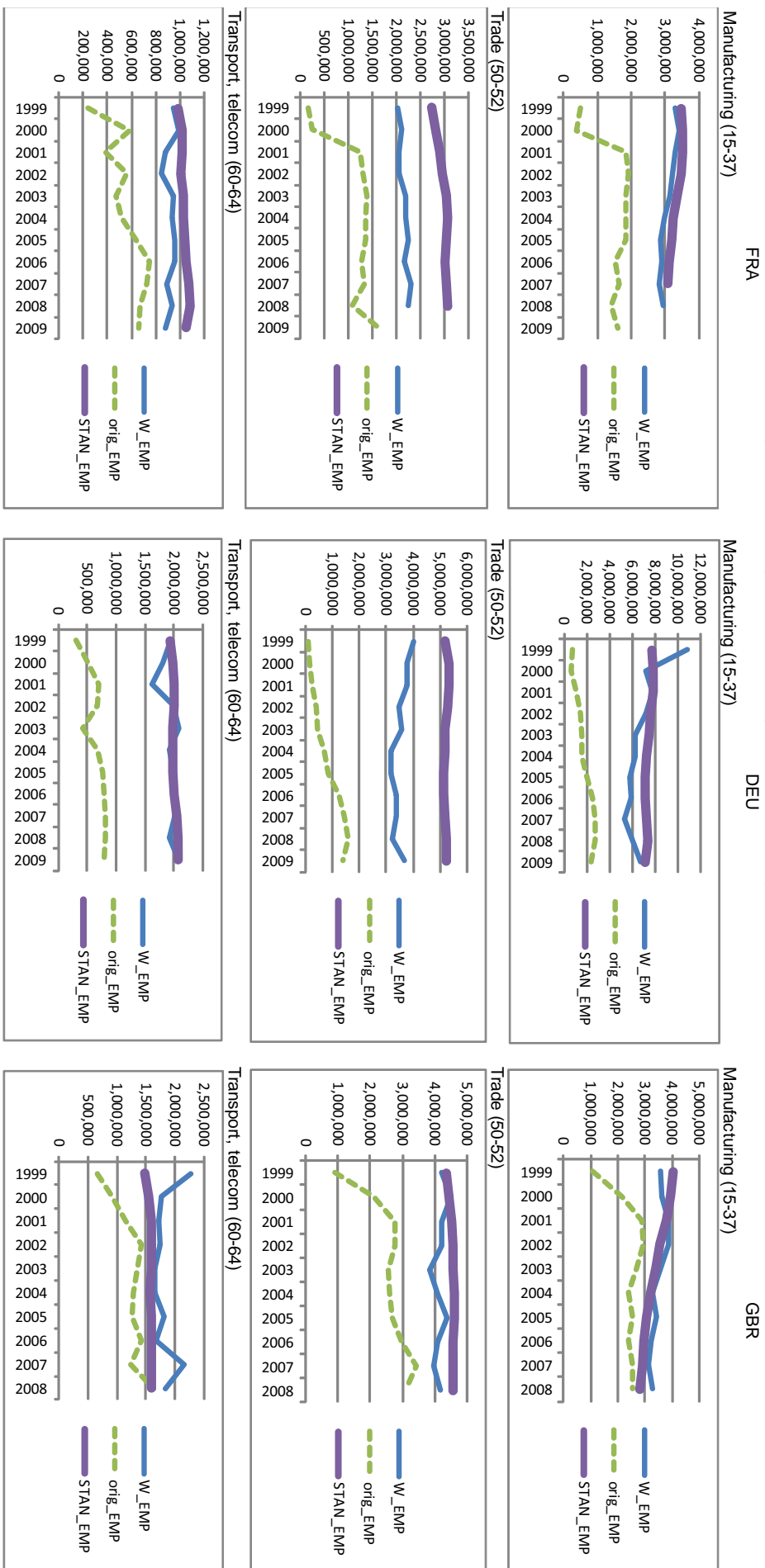
¹⁹This means that for a sample where one calculates, for instance, labour productivity, there should be different resampling weights than for a sample where an estimation based TFP is used.

Table 3: The average of resampling weights by country and sizeclass (2005)

	1-9	10-19	20-49	50-249	250+	Average
AUT	397.4	165.1	45.7	6.8	2.4	79.0
BEL	1.7	1.3	1.2	1.2	1.2	1.6
CZE	6.4	2.8	1.8	1.3	1.2	4.2
DEU	333.6	126.8	26.5	5.1	1.9	46.6
DNK	3.5	3.3	2.8	2.3	2.1	3.3
EST	1.5	1.6	1.5	1.8	2.1	1.6
ESP	2.4	1.7	1.5	1.4	1.2	2.1
FIN	1.4	1.3	1.3	1.3	1.6	1.4
FRA	2.1	2.0	1.7	1.6	1.6	2.0
GBR	46.6	24.2	6.2	1.9	1.1	18.2
FRC	16.2	2.7	1.5	1.4	1.8	7.7
HUN	95.2	25.3	11.4	5.6	2.7	39.7
ITA	17.7	6.0	2.3	1.3	1.2	9.4
JPN	91.5	22.6	8.4	3.6	1.1	15.7
KOR	36.0	7.7	3.1			11.7
NLD	101.8	49.0	10.1	3.0	1.1	40.8
NOR	39.2	48.6	63.5	67.0	19.0	41.7
POL	118.1	8.4	5.5	2.8	2.4	20.9
PRT	446.1	100.7	37.3	11.2	4.2	147.5
SWE	1.2	1.4	1.4	1.6	1.6	1.2
SVN	3.9	2.3	1.7	1.3	1.1	3.1
SVK	8.5	4.0	1.9	1.4	1.2	4.4
USA	51439.6	13690.8	1734.0	861.1	25.7	5192.4
Average	29.6	32.6	13.4	13.8	3.5	26.4

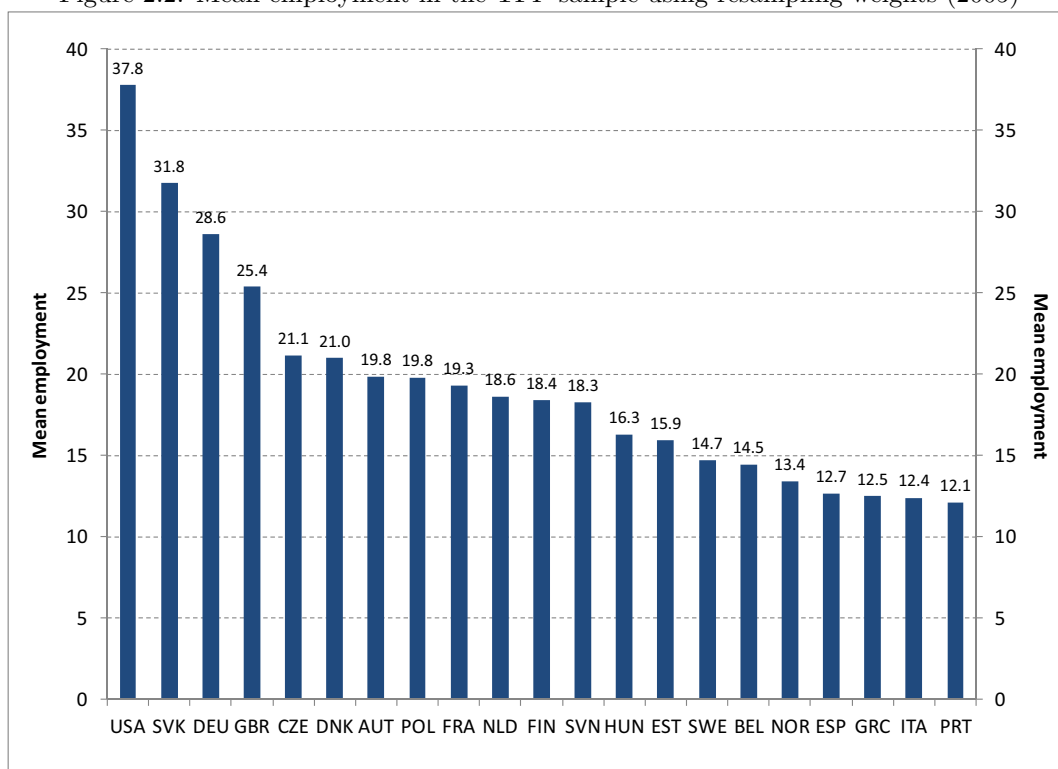
Entries show the average of the constructed resampling weights for 2005, for the Solow-residual TFP sample (using external imputations). For Korea and Japan, some industries and size classes are not contained in the SDBS database, hence no resampling weights can be constructed for those. Differences in coverage mainly reflect differences in the nature of data providers Bureau van Dijk (the vendor of ORBIS) has access to. See also Table E.5 in the Appendix.

Figure 2.1: Weighted vs. unweighted total employment in selected countries



Note: The vertical axis measures the total number of employees, in industries defined in the header of the figures (with NACE codes in parentheses). *W_EMP* denotes ORBIS employment numbers, using country*year*industry*sizeclass employment-based resampling weights, *orig_EMP* denotes ORBIS employment without any weighting but with non-missing data for TFP calculations, *STAN_EMP* denotes employment from the OECD STAN database. Missing years are due to missing data in OECD STAN or SDBS. Even though there is usually a substantial improvement from the unweighted sums, discrepancies which still persist can be due to (1) approximations used during filling up the SDBS, (2) differences between SDBS and STAN on their definition of firm coverage and (3) error due to randomness in the procedure of constructing weights.

Figure 2.2: Mean employment in the TFP sample using resampling weights (2005)



Note: Each column shows the weighted mean of employment for the sample where the simplest, Solow-residual type TFP measure (externally imputed) can be calculated, using the resampling weights described in Section 2.4. Japan and Korea are not included because the external database (SDBS) is not available for each of their sectors and firm size classes.

Finally, resampling implicitly assumes that firms in ORBIS *within* a specific country * industry * sizeclass * year cell are representative of the true population within that cell.²⁰ This is the justification for random replication of firms *within* cells. Resampling can correct for potential discrepancies between the true population and the ORBIS sample only *across* those cells. However, it cannot correct for potential selection bias concerning the firms included in ORBIS along other dimensions (e.g. age, profitability) or across more detailed size classes and industries than the ones in SDBS. Consequently, analyses related to firm demographics (entry, exit, etc.) face limitations.²¹

2.5 Deflation and Internationally Comparable Price Levels

The procedure of obtaining price indices is fairly standard, by using external information (OECD STAN) on industry level deflators. For internationally comparable price levels, the procedure is a bit more elaborate, involving industry-level PPP estimates by Timmer et al. (2007). As they explain, there are a number of reasons why industry PPPs instead of aggregate ones are preferred when using them for industry- or firm-level productivity analysis. First, price levels may differ substantially across industries; a typical pattern is a higher services-goods price ratio in more developed countries, related to the Balassa-Samuelson effect. Second, aggregate PPPs are based on an expenditure approach rather than an output approach, and the former is less appropriate for productivity analysis as it excludes exports but includes imports, while it is also affected by transport costs, taxes and subsidies. Finally, the output of

²⁰In other words, those firms which are observed in ORBIS – for instance, with less than 10 employees in retail trade (industry 51) in Germany – are representative of all firms in that cell.

²¹Where available, the incorporation date (INCORP_DATE) variable in ORBIS can be used to construct firm age and investigate firm behaviour upon entry, but young firms are underrepresented. Further, it is difficult to identify real exit.

certain industries, typically used as intermediate goods in other industries, are not captured at all in an expenditure approach. Nevertheless, while the cross-country comparability is likely to improve by using industry PPPs, comparing TFP levels across industries - which is inherently difficult due to inherent differences in the production technologies - may still require caution, as measuring industry PPPs also involves some trade-offs.²²

The database compiled by Timmer et al. (2007) is the most recent and most complete industry level PPP data, covering a wide range of countries and industries. Given that the comparison year used is relatively dated (1997), local currency deflators are applied changes to proxy more recent changes over time. To the extent that the cross-country and cross-industry differences (as cited above) are likely to be more important than the possible bias caused by not updating PPP estimates over time, this approach is more preferable to using aggregate PPPs. However, productivity analysis for the manufacturing sector may be carried out, as a robustness check, using market exchange rates, since price level differences in manufacturing are less pronounced due to the stronger exposure to international trade. For this reason, the constructed TFP database includes both PPP-adjusted and non-adjusted series.²³ See the detailed description in Section B of the Appendix.

3 Productivity Measures and the Resulting TFP-Database

Having described the preparations necessary to extend the database and enhance international comparability, this section discusses the various productivity measures which are calculated using ORBIS, presents the country-coverage and other properties of the resulting data set, and highlights some caveats to keep in mind.

Productivity Measures: Some Background

Productivity, in its broadest interpretation, is meant to capture the efficiency by which inputs are turned into outputs (Hulten, 2001). From the least to the most data-demanding methodologies, an array of productivity measures which can be calculated using the OECD-ORBIS database are considered below (see Table 4).

²²For more details, see Section B.2 of the Appendix.

²³The latter converted at the market exchange rates in the middle of our sample (2005) in order to minimize the impact of short-run exchange rate fluctuations on productivity levels.

Table 4: Productivity measures and possible imputations in ORBIS

Productivity measures		Definition	Possible imputations
Labour productivity	Based on gross output: LP_TR	$\log\left(\frac{TR}{L}\right)$	No
	Based on value added: LP_VA	$\log\left(\frac{VA}{L}\right)$	For VA
TFP: index-numbers	Solow-residual: TFP_Solow^*	$va - (1 - s_L)k - s_L l$	For VA and s_L, s_K
	Superlative index: TFP_SupIn^{**}	$\widetilde{va} - (1 - \alpha_L)\widetilde{k} - \alpha_L \widetilde{l}$	For VA and α_L
TFP: prod.fn. based	OLS: TFP_OLS^{***}	$va - \widehat{\beta}_K^{OLS} k - \widehat{\beta}_L^{OLS} l$	For VA
	OLS fixed effects: $TFP_OLS_FE^{***}$	$va - \widehat{\beta}_K^{OLS_FE} k - \widehat{\beta}_L^{OLS_FE} l$	For VA
	Levinsohn-Petrin : TFP_LevPet^{***}	$va - \widehat{\beta}_K^{LevPet} k - \widehat{\beta}_L^{LevPet} l$	For VA and II ⁺
	Wooldridge: $TFP_Wooldridge^{***}$	$va - \widehat{\beta}_K^W k - \widehat{\beta}_L^W l$	For VA and II ⁺

Note: Each measure varies over firm and time, though for simplicity, firm and time indices are not shown. Each variable with a monetary value (turnover, value added, capital) refers to real (i.e. deflated) values. Small letters (va, l, k) denote natural logarithms. For calculating the capital stock by using the standard Perpetual Inventory Method (PIM) using the book value of fixed tangible assets and depreciation, see Section A of the Appendix.

* s_L denotes the share of value added which goes to labour, and calculated using the OECD STAN database, per country and industry or only per industry (both measures are calculated), as the ratio of labour costs to value added.

** α_L denotes the average of the labour share in the reference firm (taken to be the average across firms in the same country*industry) and the current firm. For each variable X, \widetilde{X} is the difference from the current firm and the reference value, which is taken to be the industry average, over time and over firms, either per country or using cross-country average. In the ORBIS-TFP database, both measures are present, see Table F.1 in the Appendix.).

*** Estimation based measures assume that the production function parameters β_K and β_L are the same for a group of firms. In the calculations below using ORBIS, it is taken to be 2-digit industry groups (see Table E.4 for the exact breakdown).

+II stands for intermediate inputs.

Gross output or total revenue (TR) based labour productivity is the most widely available, thus serves as a benchmark for the analysis, but its most immediate problem is that it does not control for intermediate input usage.¹ In other words, a company with a lot of reselling activity (e.g. retail companies) will probably rank very high in this measure. Value added based labour productivity takes care of this problem, as value added itself is the difference between output (sales or revenue) and intermediate inputs (including resold goods, typical in retail trade).

However, labour productivity does not control for differences in capital intensity across firms. Therefore, in order to control for capital intensity, total factor productivity (TFP) should be calculated. Two broad types of TFP measures are considered: index number approaches and estimation-based methods.

Index number measures are easy to calculate: they simply relate output to a weighted sum of inputs. However, they usually assume constant returns to scale.² Two widely used index numbers are the Solow-type and the Superlative-index³ measures. The latter takes a reference value (an average across firms or a specific firm, e.g. the median) and implicitly compares productivity levels to that reference value, and more importantly, it uses firm-level factor shares. As such, it is more prone to measurement error than the Solow measure (see Biesebroeck (2007)).

Turning to the estimation based TFP-s, the simplest benchmark is the residual from OLS regressions when estimating the production function, usually run industry-by-industry, on firm-level data. However, the coefficients obtained through this exercise are likely to be inconsistent and biased. In the case of the labour coefficient, for instance, firms with higher productivity hire more workers, but productivity is not directly observed during production function estimation, hence it enters the error term. Firms' behavior will thus introduce a positive correlation between the error term and the labour input, rendering standard OLS inconsistent and biased.⁴ Fixed effect estimators can only partially solve the problem since they take into account only a time-invariant firm-specific productivity effect and cannot control for cases where productivity shocks occur differently across firms and time. A well-known remedy for this problem was proposed by the semi-parametric approach

of Olley and Pakes (1996), who took investment as a proxy variable for unobserved TFP. However, investment is proven to be a 'lumpy' variable and hence potentially a poor proxy for productivity.⁵ The method of Levinsohn and Petrin (2003) takes a similar approach, but instead of investment, it uses intermediate inputs as proxies. More recently, Wooldridge (2009) developed a technique which is built upon the Levinsohn and Petrin (2003) method, and it deals with the critique of Akerberg et al. (2006) on a technical issue, while also being a quicker, one-step procedure with consistent standard errors without bootstrapping.⁶

Overall, the theoretical advantage of the estimation based methods (no assumption of constant returns to scale needed) comes at a cost: their data requirement is usually higher, both in terms of the number and the quality of the available variables. Finally, it is interesting to note that, irrespective of the sophistication of the measures, it has been shown that firms' productivity levels and dynamics vary a lot, even in narrowly defined industries (see Bartelsman and Doms (2000); Syverson (2011)). This issue will be addressed again below in Section 4.1 for the TFP measures from ORBIS.

3.1 Details of the Resulting TFP Database

3.1.1 Country Coverage

The potential set of countries for TFP measurement is contained in Tables E.1 and E.2 in the Appendix. What constrains this set is the availability of external data on price indices (large emerging economies like Brazil, China, India, Indonesia, Russia, South Africa and Turkey do not have 2-digit price indices, the same applies for Ireland and Chile) and/or industry PPPs for cross-country productivity level comparisons (Switzerland and Norway are not included in the ICOP industry PPP database). Nevertheless, for the latter two countries, by using local currencies or using 2005 euro/local currency exchange rates, productivity can also be calculated or estimated. Finally, countries with a very low number of observations in ORBIS (Australia, Canada, Chile, Iceland, Ireland, Israel, Mexico, New Zealand) are also omitted from the current calculations. A summary of the final set of countries included in the database is shown in Table 5

In the June 2011 version of OECD-ORBIS, data for the United States and Canada covers only single-establishment firms. The Statistics Directorate of the OECD developed a procedure to aggregate the data of establishments which belong to the same firm into firm-level data, but this procedure was not carried out for variables required for value added based labour productivity calculations and TFP estimation

¹Note that if one considers within-sector patterns, this is less of a problem, since firms in the same sector tend to have similar intensities of intermediate input usage. Revenue based labour productivity has been used, for instance, in Foster et al. (2001), where the authors mention that the results in terms of their decompositions and relative productivity are very similar across productivity measures.

²Note that usually all TFP measures assume that firms are price takers on both output and input markets and maximize profits. Also note that assumptions about returns to scale are not an important empirical issue, as the approaches below score relatively well irrespective of such assumptions according to the experiments on productivity measurement by Biesebroeck (2007).

³For the original idea, see Caves et al. (1982) and for applications, see Griffith et al. (2009) and Arnold et al. (2008).

⁴Endogeneity arises because as productivity changes, optimizing firms react by adjusting their inputs (right hand side variables), and productivity also directly affects value added (the left hand side variable).

⁵Their method also controls for selection bias. Selection can cause a problem since firms with higher capital stock will more easily weather negative productivity shocks and stay in the sample than others.

⁶The Akerberg et al. (2006) critique basically refers to a collinearity problem: the joint inclusion of the nonparametric, polynomial terms of the variable input (e.g. labour) together with its structural coefficient in the production function makes the latter potentially unidentified.

Table 5: The set of OECD countries in ORBIS and in the ORBIS-TFP database

Country code	In the TFP sample	In the Labour Productivity sample	Resampling weights	Currency conversion using industry PPP-s*
AUT	yes	yes	yes	yes
AUS	no	no		
BEL	yes	yes	yes	yes
CAN	no	no		
CHE	yes	yes	yes	no
CHL	no	no		
CZE	yes	yes	yes	yes
DEU	yes	yes	yes	yes
DNK	yes	yes	yes	yes
EST	yes	yes	yes	yes
ESP	yes	yes	yes	yes
FIN	yes	yes	yes	yes
FRA	yes	yes	yes	yes
GBR	yes	yes	yes	yes
GRC	yes	yes	yes	yes
HUN	yes	yes	yes	yes
IRL	no	no		
ISR	no	no		
ISL	no	no		
ITA	yes	yes	yes	yes
JPN	yes	yes	yes ⁺	yes
KOR	yes	yes	yes ⁺	yes
LUX	no	no		
MEX	no	no		
NLD	yes	yes	yes	yes
NOR	yes	yes	yes	no
NZD	no	no		
POL	yes	yes	yes	yes
PRT	yes	yes	yes	yes
SWE	yes	yes	yes	yes
SVN	yes	yes	yes	yes
SVK	yes	yes	yes	yes
TUR	no	no		
USA	yes	yes	yes ⁺⁺	yes

⁺Japan and Korea have resampling weights for only a subset of their firms due to the lack of complete availability of the SDBS database on which resampling is based.

⁺⁺For the United States, only single establishment firms are used from the ORBIS database. See more on this issue in Section 3.1.1.

* Switzerland and Norway are not in the industry-level ICOP PPP 1997 database which is used for the calculations. They are in the database which uses 2005 actual euro-local currency exchange rates as the conversion.

Countries not included in the productivity database have very few observations where variables for TFP are available (Australia, Canada, Chile, Iceland, Ireland, Israel, Luxembourg, Mexico, New Zealand, Turkey, see Tables E.1 and E.2 in the Appendix). Large emerging non-OECD countries (Brazil, China, India, Indonesia, Russia and South Africa) are excluded mainly due to the lack of price indices, but for some, there are few observations available for productivity calculations. Switzerland has no external, industry-level wage data (labour compensation) available in OECD STAN, hence it cannot be used with the externally imputed TFP calculations.

(Tangible fixed assets and EBITDA are missing).^{24,25}

3.1.2 Sectoral Coverage

Throughout the analysis, the coverage is always the non-farm business sector (i.e. without agriculture and public services).²⁶ Productivity estimates for the financial sector (NACE Rev 1.1 codes 65-67) can be calculated in principle, but they should be treated with caution due to the inherent difficulties in measuring output.²⁷

3.1.3 Industry Breakdown

It is not obvious at which level of industry breakdown homogeneity in the production function parameters can be safely assumed. More detailed, lower levels of aggregation are desirable, but the resulting reduction in sample size during industry-by-industry production function estimation might create problems, especially for small countries. While assuming homogeneity at the 2-digit industry level is generally considered a good compromise and is feasible in most industries and countries in ORBIS, it still yields a very low number of observations for certain country-industry pairs. In such cases, the EU-KLEMS industry classification is applied (O'Mahony and Timmer, 2009)²⁸, which merges and treats together some 2-digit industries producing similar goods (e.g., in NACE Rev 1.1, Textiles (17), Wearing Apparel (18) and Leather and Footwear (19) are merged into one industry group (17-19).) See Table E.4 of the Appendix for more details.

3.1.4 Filtering for Outliers

By default, no outlier-filtering is implemented, but in principle, there are two possibilities:

1. Filtering observations before the computation / estimation, based on extreme levels of *ratios* (K/L , M/Y , etc.) and erratic *movements* (dVA , dL) of inputs or output.
2. Filtering observations after the estimation, based on extreme levels and changes of productivity.
Both methods should be done by country and industry, that is, at the aggregation level at which

the estimations are run. The first type of filtering is probably more important for estimation based approaches, whereas the second technique provides a useful crosscheck for the results in any case. In the current version of the database, no such filtering is implemented, but the second type can be easily carried out using the available productivity estimates.²⁹

3.1.5 Further Criteria for TFP Estimates

The following criteria is applied when using TFP estimates based on production functions:

²⁴As such, in the future, there may be scope for including additional, multi-establishment firms of these countries in the productivity calculations if the procedure is extended to these variables.

²⁵The number of observations for each country, year and imputation type, for the Superlative Index measure, are contained in the *TFP_Sample.xlsx* file in the root directory of the TFP database.

²⁶That is, industries 15-74 in the NACE industry classification Rev 1.1. Mining and quarrying is also excluded due to the low number of firms present in those industries in many countries.

²⁷Note that to a lesser extent, this is also true for other business services, where output measurement is more problematic than in manufacturing sectors.

²⁸In the core database, estimation based measures are calculated at the EU-KLEMS 2-digit industry group levels, and the index number approaches are calculated at the 2-digit industry levels.

²⁹There is one exception though: as the Superlative Index measures seem particularly sensitive to imputations, in line with the findings in Biesebroeck (2007) about its sensitivity to measurement errors, a type of filtering method called *Winsorisation* is applied when constructing the expenditure shares needed for the calculations. This means that values in the extremes, i.e. outside some pre-defined top and bottom percentiles, are replaced by those top and bottom percentile values. In the current situation, the labour cost shares in the top and bottom 1 percent of observations are replaced by the 1st or the 99th percentile values, for extreme small and extreme large values, respectively, per industry (and per country, in the case of country-by-country reference values). This is a practice followed by recent studies using firm-level data (e.g. Sivadasan (2009); Gabaix (2011)).

1. Within each country * industry group, at least 50 observations are required for the estimation.
2. Both the labour and capital coefficients need to be larger than zero.

In the few cases where these do not hold (typically, for small countries and with the more detailed industry breakdowns), values of estimation based TFP measures are left empty.³⁰

3.2 Qualifications and Limitations with TFP Measures from ORBIS

The strength of the data set is that it allows users to compare many countries and industries throughout more than a decade, using standardized methodologies for firm-level TFP estimations. Nevertheless, a number of caveats should be kept in mind:

- Regarding the measurement of output, as in most other firm-level data sets, ORBIS does not contain firm-level price indices, hence only industry-level deflators can be used, following the best practice in the literature (see Section B of the Appendix for details). The disadvantage is that measured productivity differences within the same industry may also reflect differences in market power embodied in prices (Syverson, 2011).
- Regarding labour input, the database does not contain the number of hours worked, only the number of employees. It is also not possible to consider different types of workers (i.e. low skilled or high skilled) as there are no information on the characteristics of employees. This could in principle be circumvented by assuming that total labour costs are a measure of quality and intensity adjusted labour, but it is not always available.³¹ Also, total labour costs are directly influenced by changes in the regulatory environment (social security contributions, minimum wages, etc.). For this reason, only the number of employees is used as labour input, thus the productivity measures should be interpreted as also capturing labour quality and intensity (i.e. capacity utilization).
- The type of capital goods in ORBIS are differentiated only to the extent of being tangible and intangible, but there is no differentiation regarding the type of the asset (e.g. structures or equipment). In order to avoid conceptual difficulties in measuring and valuing intangibles, which are poorly reported in ORBIS, only data on total tangible fixed assets are used. Further, since only the stocks but not actual capital services are observed, measured productivity changes may capture changes in capacity utilization, similarly to the lack of available hours worked information in the case of labour input.
- Intermediate inputs are not differentiated across materials, energy and purchased services, only materials can be observed.³² The lack of availability of detailed intermediate inputs prevents the calculation of output based TFP-s using a more detailed production function (see final point below).
- Finally, in the resulting TFP-database, only value added based TFP measures are considered in order to minimise the potential problems related to the measurement of intermediate inputs, and output based measures are left out.³³

³⁰These results are available on request. Coefficients for further inspection are stored next to the database in separate files, see Section F of the Appendix on the structure of the database.

³¹In fact, the lack of availability of labour costs in some countries (notably, the United States) necessitates the use of external imputations, see Section 2.3.

³²Alternatively, intermediate inputs can be calculated as the difference between output (measured by turnover) and value added. This is the method used in the specifications with internal imputations.

³³Output based measures estimate the production function as having production as output and capital, labour and intermediate input (preferably broken down by energy, purchased services and materials) as inputs. As such, they can better control for differences in the production structure across industries, but they require more information, and their aggregation is more complicated than those of value added based measures. On the relationship between the two measures, and on the unsettled issues regarding their relative merit, see the detailed description in OECD (2001).

4 Comparison of Productivity Measures

This section explores two main issues for countries with either reasonably good data availability (Spain, Italy and Great Britain, in that order), or medium level data availability (France, Germany): (1) how well the productivity measures based on imputed data approximate the ones based on actual data, in terms of levels as well as growth rates, (2) do we get a distorted picture of within-industry productivity distributions when using imputations? Indeed, it is essential to check whether TFP measures using imputations still preserve a large part of the within-industry heterogeneity, since imputing firm-level values for the labour cost component of value added by relying partly on industry-level data can reduce the within-industry variation, both in value added and TFP.

The results of these analyses show that using internally imputed values give a very good approximation of actual values. Hence such imputations can be applied to increase sample size. External imputations, although less accurately approximating actual values, also work reasonably well.³⁴

The relationship among productivity measures using different methodologies is also investigated. The measures considered are the Solow-residual and the superlative index, which are index-number based, and estimated measures based on OLS-residuals and on the recent semi-parametric approach by Wooldridge (2009). This analysis relies exclusively on non-imputed data and can potentially shed light on the extent to which the underlying assumptions behind these different methodologies result in different TFP values.

4.1 Quality of Imputations for TFP Measures

4.1.1 Correlations

Table 6 shows correlations of TFP measures based on imputations with those based on actual, non-imputed data, on the subset of firms where both can be calculated. As such, each entry in the table shows

$$\text{Corr} \left(\ln \left(TFP_{it}^{\text{non-imputed}} \right), \ln \left(TFP_{it}^{\text{imputed}} \right) \right) \quad (4.1)$$

for the left panel, and

$$\text{Corr} \left(\Delta \ln \left(TFP_{it}^{\text{non-imputed}} \right), \Delta \ln \left(TFP_{it}^{\text{imputed}} \right) \right) \quad (4.2)$$

for the right panel, where subscript i refers to firms, t refers to years, and Δ is the first difference operator. In order to isolate the time dimension, these correlations are computed for a particular year. In Table 6, it is $t=2008$.³⁵ A number of key findings emerge:

- First, the variation in the internal and the refined external imputations is very similar to the variation in the non-imputed values, with correlations around 0.8-0.9 for all types of productivity measures and for levels, and around 0.6-0.85 for growth rates, in the average of the five countries considered in the comparison. The refined external imputations perform consistently better than the non-refined ones, making the case for predicting wage differentials from the industry average (see equation 2.4).³⁶

³⁴Given the potential for more serious measurement problems with small firms' data for each of these exercises, estimates are constructed and comparisons are carried out both for all firms and for the subset of firms with 20 or more employees. On the structure of the resulting ORBIS-TFP database, see the description in Section F of the Appendix.

³⁵Correlations for other years show very similar patterns and are available upon request.

³⁶Note that for countries where labour costs are not observed at all (United States and Greece), predictions are based on regressions of within-industry wage differentials on firm characteristics in countries with similar labour market characteristics (for the United States, Great Britain is used, and for Greece, Spain and Portugal are used). These predictions hence may be less accurate than for the countries presented in this section. Nevertheless, correlations with external imputations *not* involving wage predictions serve as a lower bound for the correlation between the refined externally imputed values and the true unobserved values also for the United States and Greece, which stand at acceptable 0.7-0.8 for TFP levels and 0.6 for growth rates, the Superlative Index performing a bit worse than the others.

- Second, external imputations yield the strongest relationship with non-imputed values for the Solow-type measure, both in levels and growth rates: on average, around 0.9 and 0.7, respectively. The Superlative Index measure is somewhat less well captured (0.8, 0.7). Among estimation type measures, OLS gives a slightly closer match than Wooldridge, for which correlations between the growth rate show the lowest values (0.6 on average).
- These patterns are true for the subsample of firms with at least 20 employees as well as for the whole economy, or using simple correlations versus rank correlations. When using rank correlations, imputations for the Superlative Index measure tends to show a bit stronger relationship with the non-imputed values than just looking at simple correlations (see Table E.6 in Appendix). In addition, even though Table 6 is based on the whole market sector and for the year 2008, the patterns are preserved also when the comparisons are made sector-by-sector (Table 7) and in other years.

Table 6: The effect of imputations on productivity measures: correlation between non-imputed, internally and externally imputed measures, for selected countries

ESP									
Correlations for ln(TFP)					Correlations for TFP growth				
<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge	<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge
<i>imputation types</i>					<i>imputation types</i>				
internal	0.98	0.90	0.97	0.97	internal	0.97	0.92	0.95	0.95
external	0.85	0.54	0.75	0.78	external	0.81	0.69	0.74	0.71
external, refined	0.92	0.77	0.84	0.85	external, refined	0.88	0.82	0.78	0.71
ITA									
Correlations for ln(TFP)					Correlations for TFP growth				
<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge	<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge
<i>imputation types</i>					<i>imputation types</i>				
internal	0.98	0.91	0.96	0.96	internal	0.90	0.87	0.89	0.87
external	0.86	0.66	0.72	0.73	external	0.72	0.64	0.64	0.56
external, refined	0.92	0.82	0.82	0.80	external, refined	0.80	0.73	0.71	0.60
GBR *									
Correlations for ln(TFP)					Correlations for TFP growth				
<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge	<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge
<i>imputation types</i>					<i>imputation types</i>				
internal	0.96	0.76	0.92		internal	0.84	0.79	0.81	
external	0.82	0.52	0.70		external	0.64	0.49	0.58	
external, refined	0.90	0.76	0.86		external, refined	0.69	0.67	0.61	
FRA									
Correlations for ln(TFP)					Correlations for TFP growth				
<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge	<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge
<i>imputation typFR</i>					<i>imputation typFR</i>				
internal	0.93	0.81	0.90	0.88	internal	0.78	0.73	0.76	0.74
external	0.74	0.52	0.62	0.68	external	0.52	0.46	0.48	0.44
external, refined	0.84	0.68	0.76	0.78	external, refined	0.60	0.53	0.52	0.45
DEU									
Correlations for ln(TFP)					Correlations for TFP growth				
<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge	<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge
<i>imputation types</i>					<i>imputation types</i>				
internal	0.95	0.81	0.90	0.89	internal	0.81	0.77	0.77	0.76
external	0.88	0.71	0.67	0.72	external	0.60	0.59	0.58	0.56
external, refined	0.94	0.84	0.86	0.79	external, refined	0.69	0.68	0.64	0.61
Average over ESP, ITA, GBR, FRA, DEU									
Correlations for ln(TFP)					Correlations for TFP growth				
<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge	<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge
<i>imputation types</i>					<i>imputation types</i>				
internal	0.96	0.84	0.93	0.92	internal	0.86	0.82	0.84	0.83
external	0.83	0.59	0.69	0.73	external	0.66	0.57	0.60	0.56
external, refined	0.90	0.78	0.83	0.81	external, refined	0.73	0.69	0.65	0.59

Note: correlations are based for observations in the year 2008, according to equations 4.1 and 4.2, for the non-farm business sector. TFP growth is defined as the first differences of $\ln(TFP)$. For a description of the imputation types, see Section 2.3.

*For Great Britain, non-imputed Wooldridge type estimates are not available due to the lack of the availability of proxies for intermediate inputs.

Table 7: The effect of imputations on productivity measures: correlation between non-imputed, internally and externally imputed measures, by sectors

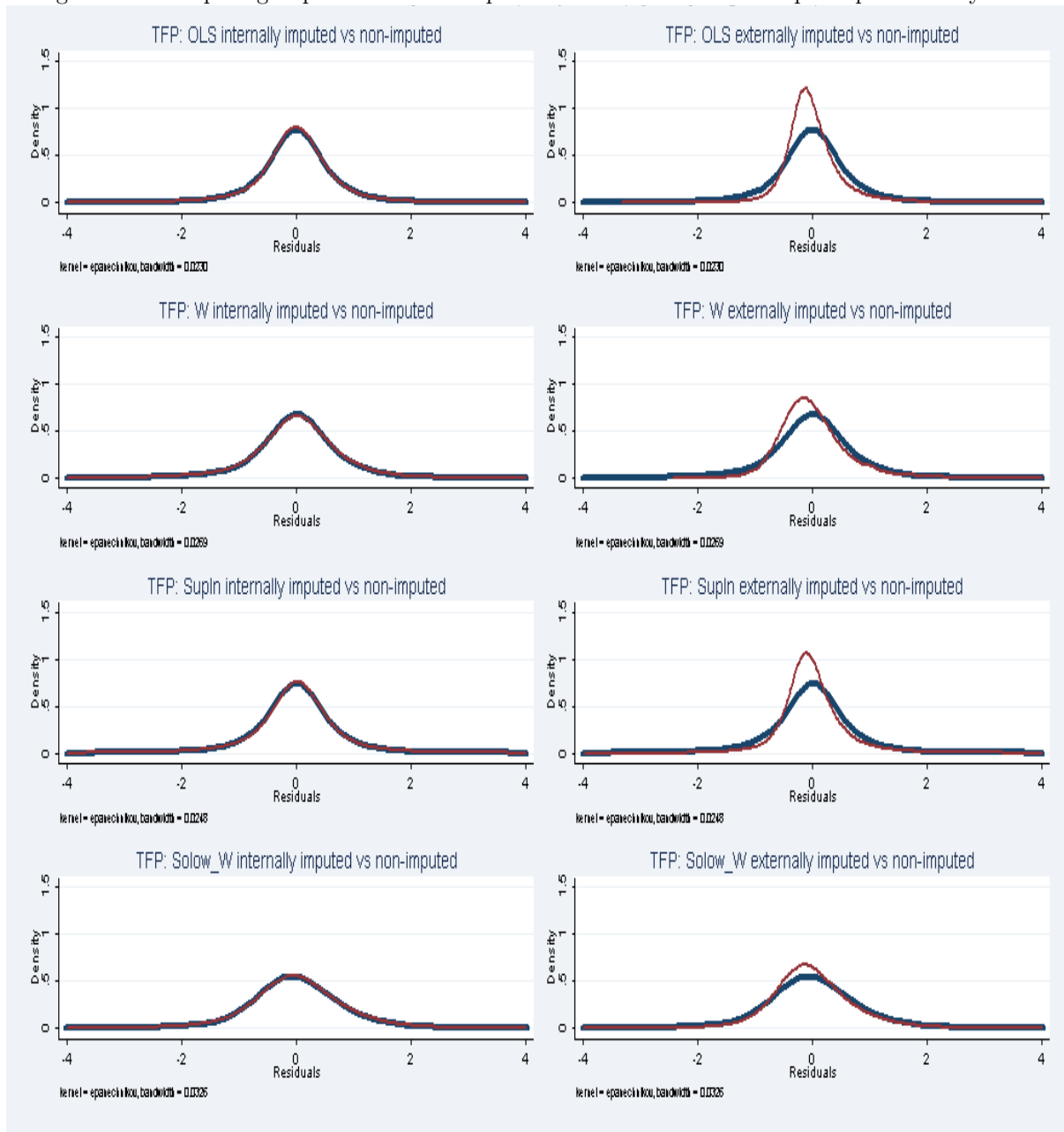
Manufacturing									
Correlations for $\ln(TFP)$					Correlations for TFP growth				
<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge	<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge
<i>imputation types</i>					<i>imputation types</i>				
internal	0.96	0.85	0.94	0.94	internal	0.85	0.82	0.84	0.84
external	0.82	0.59	0.77	0.78	external	0.63	0.58	0.61	0.58
external, refined	0.89	0.73	0.86	0.82	external, refined	0.68	0.64	0.64	0.58
Construction									
Correlations for $\ln(TFP)$					Correlations for TFP growth				
<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge	<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge
<i>imputation types</i>					<i>imputation types</i>				
internal	0.97	0.93	0.95	0.96	internal	0.91	0.90	0.90	0.91
external	0.79	0.62	0.68	0.73	external	0.70	0.63	0.66	0.60
external, refined	0.86	0.73	0.76	0.79	external, refined	0.78	0.69	0.70	0.60
Trade									
Correlations for $\ln(TFP)$					Correlations for TFP growth				
<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge	<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge
<i>imputation types</i>					<i>imputation types</i>				
internal	0.96	0.89	0.95	0.90	internal	0.87	0.84	0.85	0.85
external	0.81	0.63	0.73	0.62	external	0.67	0.61	0.63	0.61
external, refined	0.90	0.79	0.84	0.77	external, refined	0.76	0.70	0.68	0.70
Other business services									
Correlations for $\ln(TFP)$					Correlations for TFP growth				
<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge	<i>TFP measures</i>	Solow	Sup.In.	OLS	Wooldridge
<i>imputation types</i>					<i>imputation types</i>				
internal	0.95	0.81	0.89	0.89	internal	0.84	0.77	0.80	0.78
external	0.83	0.58	0.66	0.68	external	0.67	0.55	0.60	0.54
external, refined	0.91	0.80	0.78	0.78	external, refined	0.76	0.71	0.66	0.59

Note: correlations are based for observations in the year 2008, according to equations 4.1 and 4.2. Each value is an average of the values obtained in the five selected countries (Spain, Italy and Great Britain, France, Germany). TFP growth is defined as the first differences of $\ln(TFP)$. For a description of the imputation types, see Section 2.3.

4.1.2 Distributions

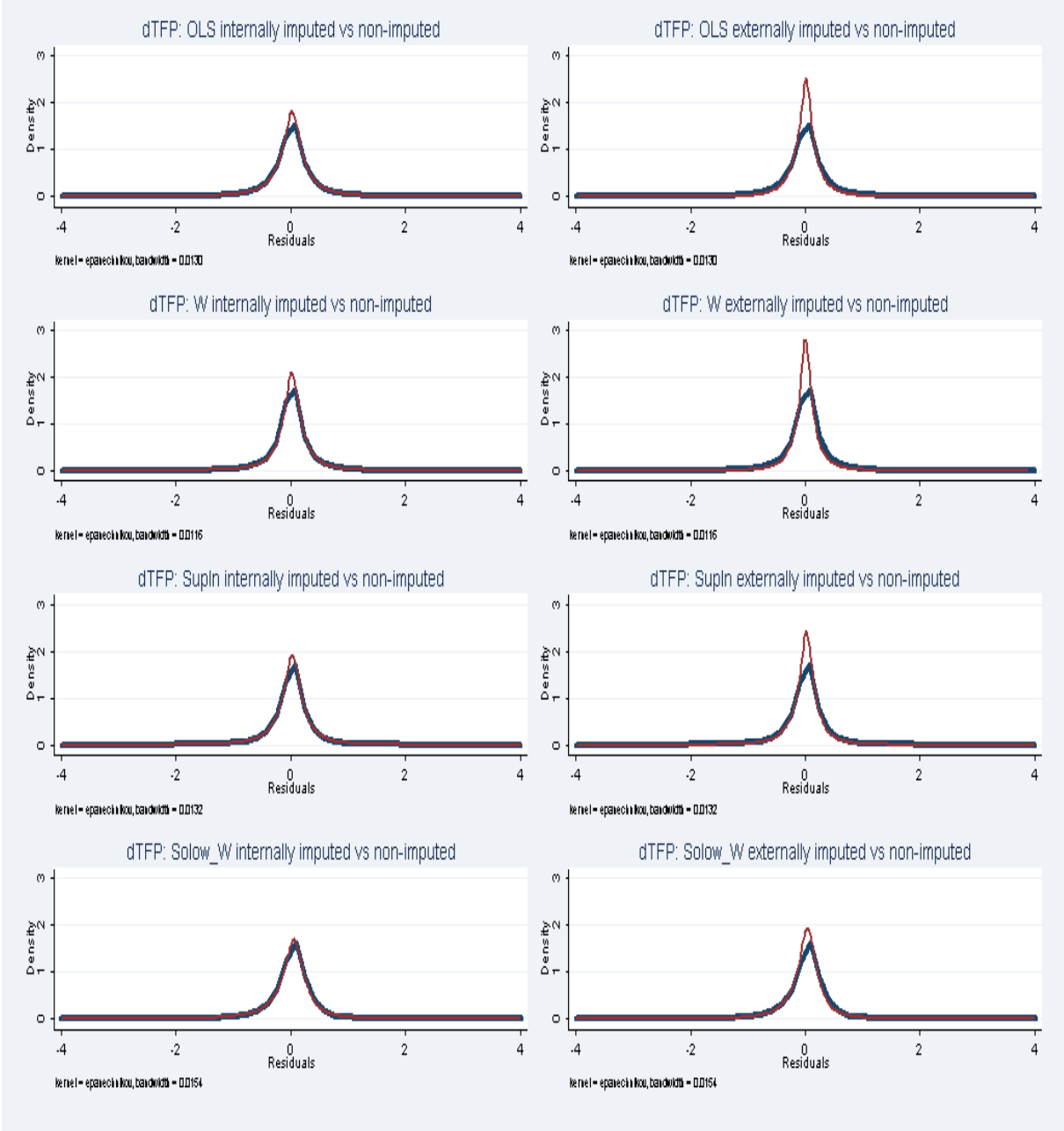
Having established that the correlations are high between imputed and non-imputed productivity measures, it is also interesting to check whether they give the same overall shape of cross-sectional distribution. Figures 4.1 and 4.2 compare the within-industry distribution of firm-level TFP-s using non-imputed and imputed values for Spain, for levels and growth rates, respectively, for four types of productivity measures. Internal imputations present a nearly identical picture for productivity levels, irrespective of the measures used. External imputations perform the best for the Wooldridge and the Solow type measures. For growth rates, the relationship is less tight, there is clearly less dispersion in the imputed values (i.e. there are more small changes). Nevertheless, a large amount of the variation is well preserved. Further, in order to see more precisely whether lower dispersion may interfere with cross-country, cross-industry analyses, the next section analyses this issue more closely.

Figure 4.1: Comparing imputed and non-imputed TFP distributions for Spain: productivity levels



Note: Figures show kernel densities of within-industry and year $\ln(TFP)$ distributions, that is, after controlling for industry * year effects. *OLS* and *W* stand for OLS and Wooldridge production function estimation based measures, *SupIn* and *Solow_W* stands for Superlative index and Solow-residual type measures. The thick lines indicate the values using no imputations, and the thin ones indicate the ones based on imputations (left panel: internal, right panel: refined external). For each productivity measure, the sample where the densities are shown is the joint sample where both imputed and non-imputed values are available.

Figure 4.2: Comparing imputed and non-imputed TFP distributions for Spain: productivity growth



Note: Figures show kernel densities of within-industry and year $\Delta \ln(TFP)$ distributions, that is, firm-level TFP growth, after controlling for industry * year effects. OLS and W stand for OLS and Wooldridge production function estimation based measures, $SupIn$ and $Solow_W$ stands for Superlative index and Solow-residual type measures. The thick lines indicate the values using no imputations, and the thin ones indicate the ones based on imputations (left panel: internal, right panel: refined external). For each productivity measure, the sample where the densities are shown is the joint sample where both imputed and non-imputed values are available.

4.1.3 Dispersion by Industry

As noted above in the Box on Productivity Measures, a key stylized fact from firm-level empirical studies is that productivity is widely dispersed even within narrowly defined industries. A large observed dispersion may indicate that there is a large room for aggregate productivity increases by re-allocating resources from the less to the more productive firms (see Hsieh and Klenow, 2009). At the same time, a wider range of outcomes may indicate a greater degree of experimentation and learning by doing (see Bartelsman et al., 2010). For these reasons, it is important to check how differences in dispersion are preserved when TFP measures based on imputed data are used.

Table 8 summarizes some key statistics on dispersion for selected countries and for the measures based on production function estimations, comparing imputed and non-imputed measures. Dispersion is calculated as the difference between the 90th and the 10th percentile of log TFP (see Syverson, 2011). The key lesson from the results based on our five selected countries in Table 8 is that dispersion levels are lower for the externally imputed values (on average, 0.99 for OLS, 1.45 for Wooldridge) than for the non-imputed (OLS: 1.31, Wooldridge: 1.55) or internally imputed values (OLS: 1.35, Wooldridge: 1.74). This is clearly driven by the fact that part of the within-industry dispersion in productivity is not captured when using external values for the average labour costs. The pattern is similar across all five selected countries, and the differences are in the range of 0.3-0.4 log points for each country for OLS measures and 0.2-0.3 for the Wooldridge measure. For illustrative purposes, the dispersion levels by industry are shown for Spain on Figure 4.3.³⁷ Note that the range of values is plausible in light of the reported values in Syverson (2004) (in the range of 0.65-1), using business register data for the United States and finer (four-digit) industry aggregation than what is done here (2-digit or higher).

However, the correlation across industries between the dispersions of the imputed and non-imputed values are quite high (around 0.9, see again Table 8, the lines with *Corr.*). Such high values mean that the relative differences in dispersion remain very well preserved across industries.³⁸ This result is potentially important for researchers interested in using ORBIS to explore issues related to firm-level heterogeneity.

³⁷Looking at Figure 4.3 for Spain, it is also interesting to note the differences across industries in the productivity dispersions (the patterns are quite similar in the other countries as well). First, industries with the highest productivity dispersion are energy (water and gas) transport and management (industries 40 and 41), telecommunications (64) and other business services (70-74). It is not clear what may stand behind such large dispersions of those sectors. Second, industries with the lowest dispersions are usually characterised by a high degree of homogeneity (wood processing, 20; metallic products, 27-28) or relatively strong competition (construction, 45; hotels and catering, 55). These results may warrant further exploration, as the literature is not quite settled on the reasons behind and implications of within industry TFP-dispersions.

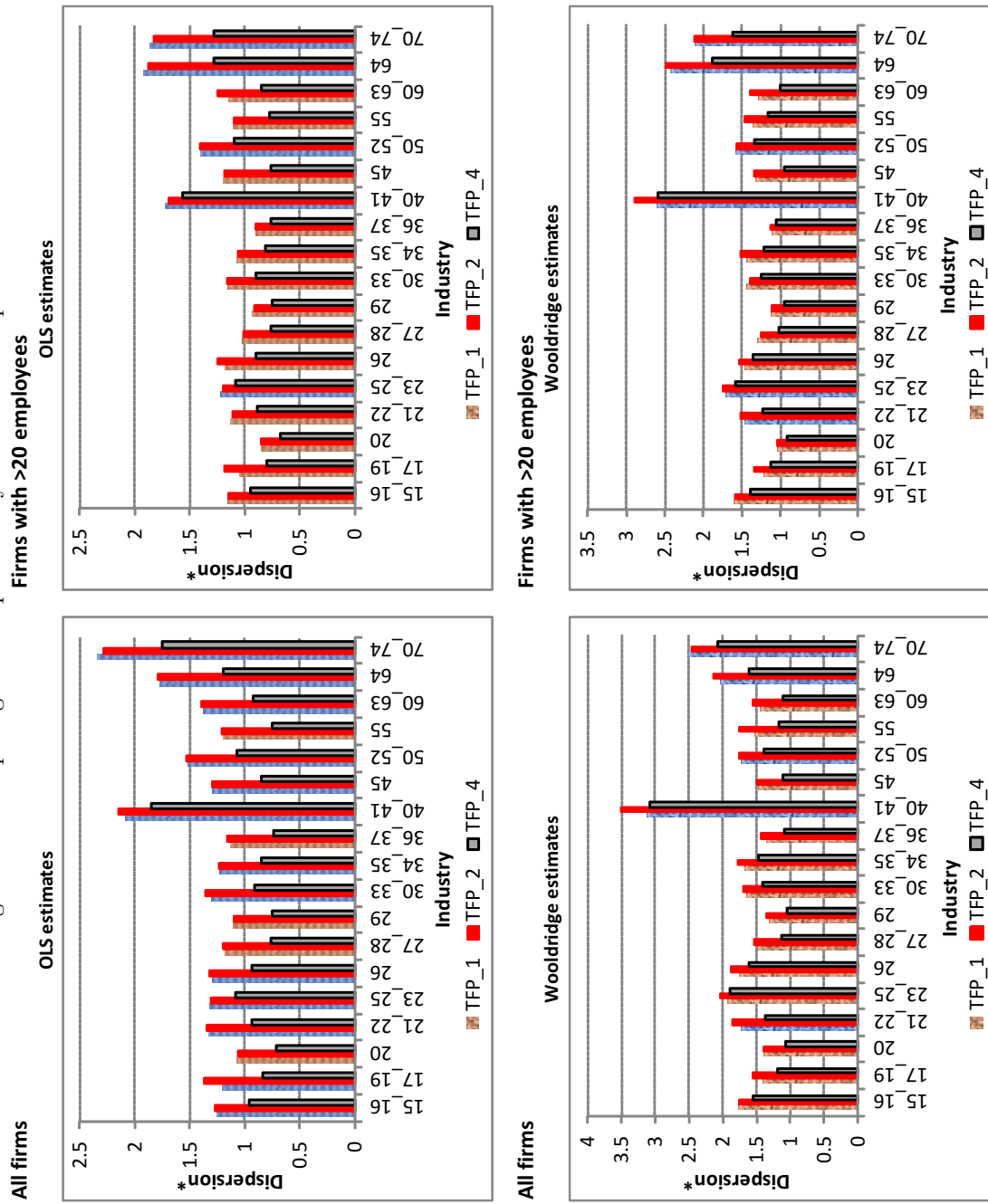
³⁸This pattern holds also for the restricted sample which contains only firms with at least 20 employees, although correlations are somewhat lower.

Table 8: Comparing within-industry TFP dispersions across imputed and non-imputed measures and for selected countries

	ESP				ITA				GBR*			
	all empl. >20 empl.				all empl. >20 empl.				all empl. >20 empl.			
Mean of industry-level dispersion measures (OLS TFP)	TFP_OLS_1	1.39	1.22	TFP_OLS_1	1.33	1.24	TFP_OLS_1	1.36	1.19			
	TFP_OLS_2	1.41	1.23	TFP_OLS_2	1.35	1.24	TFP_OLS_2	1.45	1.22			
	TFP_OLS_4	0.99	0.94	TFP_OLS_4	1.07	0.99	TFP_OLS_4	1.04	0.89			
	Corr. (1,,2)	0.99	0.99	Corr. (1,,2)	1.00	1.00	Corr. (1,,2)	0.98	0.99			
dispersion between imputations	Corr. (1,4)	0.96	0.91	Corr. (1,4)	0.98	0.88	Corr. (1,4)	0.96	0.91			
	TFP_W_1	1.73	1.53	TFP_W_1	1.45	1.37	TFP_W_1	1.95	1.50			
Mean of industry-level dispersion measures (Wooldridge TFP)	TFP_W_2	1.83	1.58	TFP_W_2	1.55	1.38	TFP_W_2	1.95	1.50			
	TFP_W_4	1.47	1.32	TFP_W_4	1.41	1.26	TFP_W_4	1.61	1.26			
Correlation of industry-level dispersion between imputations	Corr. (1,2)	0.98	0.90	Corr. (1,2)	0.99	0.96	Corr. (1,2)	0.98	0.90			
	Corr. (1,4)	0.97	0.80	Corr. (1,4)	0.98	0.81	Corr. (1,4)	0.98	0.81			
FRA												
all empl. >20 empl.												
Mean of industry-level dispersion measures (OLS TFP)	TFP_OLS_1	1.19	1.03	TFP_OLS_1	1.29	1.17	TFP_OLS_1	1.31	1.17			
	TFP_OLS_2	1.28	1.02	TFP_OLS_2	1.25	1.15	TFP_OLS_2	1.35	1.17			
	TFP_OLS_4	0.81	0.78	TFP_OLS_4	1.02	0.93	TFP_OLS_4	0.99	0.90			
	Corr. (1,2)	0.99	1.00	Corr. (1,2)	0.93	0.94	Corr. (1,2)	0.98	0.98			
dispersion between imputations	Corr. (1,4)	0.90	0.92	Corr. (1,4)	0.85	0.89	Corr. (1,4)	0.93	0.90			
	TFP_W_1	1.41	1.19	TFP_W_1	1.60	1.40	TFP_W_1	1.55	1.37			
Mean of industry-level dispersion measures (Wooldridge TFP)	TFP_W_2	1.70	1.28	TFP_W_2	1.66	1.46	TFP_W_2	1.74	1.44			
	TFP_W_4	1.23	1.05	TFP_W_4	1.52	1.30	TFP_W_4	1.45	1.24			
Correlation of industry-level dispersion between imputations	Corr. (1,2)	0.96	0.93	Corr. (1,2)	0.95	0.81	Corr. (1,2)	0.97	0.90			
	Corr. (1,4)	0.85	0.89	Corr. (1,4)	0.83	0.73	Corr. (1,4)	0.91	0.81			
DEU												
all empl. >20 empl.												
Average over 5 countries												
all empl. >20 empl.												

*GBR has no non-imputed measures for the Wooldridge (2009) type estimator because proxies for intermediate inputs are not available. Note: Dispersion is calculated as the difference between the 90th and the 10th percentile of log TFP. TFP_{OLS} refers to TFP measures as residuals from OLS regressions, TFP_W refers to TFP measures constructed by the Wooldridge (2009) semi-parametric estimator. TFP_{I_1} , TFP_{I_2} , TFP_{I_4} denote non-imputed, internally imputed and externally imputed (refined variant) TFP measures.

Figure 4.3: Comparing TFP dispersions by industries for Spain



*Dispersion is calculated as the difference between the 90th and the 10th percentile.

Note: *TFP_1*, *TFP_2*, *TFP_4* denote non-imputed, internally imputed and externally imputed (refined variant) TFP measures. The right panel uses firms with at least 20 employees (on average, over the firm's observed years). The industry groups are taken from the EU-KLEMS database O'Mahony and Timmer (2009). For the details on the industry groups, see Table E.4 in the Appendix.

4.2 Comparing Different Methodologies

From a methodological point of view, it is also interesting to check how the productivity numbers differ when comparing estimation based vs. index number approaches for TFP and how these measures relate to labour productivity. For such comparisons, we always use measures which are based on readily available (i.e. non-imputed) data. Table 9 summarizes the results.

The correlations between labour productivity (value added based) and the TFP estimates are usually quite high, both in terms of levels and growth rate, as averages across the five selected countries stand around 0.9 (for levels) and 0.95 (for growth rates). Correlations between labour productivity and TFP are usually lower and vary more across countries if we consider index-number type TFP measures (0.5, 0.8), and are usually higher for growth rates than for levels.

Turning to the relationship among TFP measures, correlations for growth rates tend to be higher than for levels, reaching nearly one for the growth rates between OLS and Wooldridge estimates. For levels, the weakest correlation is observed between the Solow residual and the Superlative Index measure, but that is also reasonably high (around 0.75).

Overall, the results indicate that different types of productivity measures, and especially their growth rates, are generally strongly related. The relationship between estimation-based measures is especially strong.³⁹

³⁹On the relationship between the OLS and Wooldridge production function estimates, see Section D of the Appendix.

Table 9: Comparing productivity measures using different methodologies
Correlations based on measures using only non-imputed variables

ESP											
Correlations for ln(TFP)						Correlations (TFP growth)					
	VA	Solow	SupIn	OLS	Wooldridge		VA	Solow	SupIn	OLS	Wooldridge
VA	1.00					VA	1.00				
Solow	0.49	1.00				Solow	0.73	1.00			
SupIn	0.56	0.70	1.00			SupIn	0.76	0.87	1.00		
OLS	0.90	0.76	0.72	1.00		OLS	0.94	0.88	0.88	1.00	
Wooldridge	0.93	0.60	0.64	0.91	1.00	Wooldridge	0.97	0.80	0.82	0.98	1.00
ITA											
Correlations for ln(TFP)						Correlations (TFP growth)					
	VA	Solow	SupIn	OLS	Wooldridge		VA	Solow	SupIn	OLS	Wooldridge
VA	1.00					VA	1.00				
Solow	0.41	1.00				Solow	0.70	1.00			
SupIn	0.54	0.75	1.00			SupIn	0.81	0.89	1.00		
OLS	0.90	0.71	0.73	1.00		OLS	0.94	0.84	0.92	1.00	
Wooldridge	0.94	0.55	0.63	0.95	1.00	Wooldridge	0.96	0.78	0.89	0.99	1.00
GBR *											
Correlations for ln(TFP)						Correlations (TFP growth)					
	VA	Solow	SupIn	OLS	Wooldridge		VA	Solow	SupIn	OLS	Wooldridge
VA	1.00					VA	1.00				
Solow	0.69	1.00				Solow	0.90	1.00			
SupIn	0.48	0.73	1.00			SupIn	0.88	0.94	1.00		
OLS	0.86	0.83	0.66	1.00		OLS	0.96	0.96	0.94	1.00	
Wooldridge						Wooldridge					
FRA											
Correlations for ln(TFP)						Correlations (TFP growth)					
	VA	Solow	SupIn	OLS	Wooldridge		VA	Solow	SupIn	OLS	Wooldridge
VA	1.00					VA	1.00				
Solow	0.67	1.00				Solow	0.87	1.00			
SupIn	0.68	0.78	1.00			SupIn	0.89	0.92	1.00		
OLS	0.90	0.86	0.79	1.00		OLS	0.95	0.93	0.96	1.00	
Wooldridge	0.89	0.67	0.70	0.90	1.00	Wooldridge	0.96	0.89	0.94	0.99	1.00
DEU											
Correlations for ln(TFP)						Correlations (TFP growth)					
	VA	Solow	SupIn	OLS	Wooldridge		VA	Solow	SupIn	OLS	Wooldridge
VA	1.00					VA	1.00				
Solow	0.28	1.00				Solow	0.86	1.00			
SupIn	0.50	0.75	1.00			SupIn	0.86	0.94	1.00		
OLS	0.80	0.55	0.65	1.00		OLS	0.94	0.93	0.95	1.00	
Wooldridge	0.74	0.21	0.45	0.82	1.00	Wooldridge	0.95	0.89	0.92	0.99	1.00
Average over ESP, ITA, GBR, FRA, DEU											
Correlations for ln(TFP)						Correlations (TFP growth)					
	VA	Solow	SupIn	OLS	Wooldridge		VA	Solow	SupIn	OLS	Wooldridge
VA	1.00					VA	1.00				
Solow	0.51	1.00				Solow	0.81	1.00			
SupIn	0.55	0.74	1.00			SupIn	0.84	0.91	1.00		
OLS	0.87	0.74	0.71	1.00		OLS	0.95	0.91	0.93	1.00	
Wooldridge	0.88	0.51	0.60	0.90	1.00	Wooldridge	0.96	0.84	0.89	0.99	1.00

Note: VA refers to value added based labour productivity, Solow refers to the Solow residual, SupIn refers to the SuperlativeIndex, OLS refers to the residual from an OLS production function, Wooldridge refers to the semi-parametric estimation method developed by Wooldridge (2009). Correlations are based for observations in the year 2008 (the pattern is similar in other years). TFP growth is defined as the first differences of $\ln(TFP)$. For Great Britain, non-imputed Wooldridge type estimates are not available due to the lack of the availability of proxies for intermediate inputs.

5 Empirical Illustrations

To illustrate the potential use of the cross-country firm-level TFP database, this section presents some novel results which exploit the increased cross-country coverage, the enhanced representativeness and price-level corrections described in the previous sections. In particular, it describes the globally most productive firms, investigates the extent of productivity convergence, and examines the link between product market regulations and productivity growth. For further illustrations and uses, see the recent work of Andrews and Cingano (2012), which utilises the resampling weights as they are described above, Andrews et al. (2013) and Andrews and Criscuolo (2013) who investigate the role of patenting on productivity and the use of labor and capital.

5.1 Which Countries Make it to the Global Frontier?

To see if the cross-country differences in productivity levels are plausible, Table 10 shows the fraction of firms in the global top 10% productivity level by country.⁴⁰ For example, in 2008, on average across industries, 35.5% of firms in the United States and 27.3% of Sweden are in the global top 10%. As expected, firms in other EU countries (Finland, France, the Netherlands, Belgium, Denmark, Germany and Great Britain) are also overrepresented in the top 10%, while firms in Spain, Italy and Portugal are less represented. At the other end, Central-Eastern European countries and Greece are clearly underrepresented.⁴¹

Changes over time, from 2000 to 2008, also seem plausible. Countries which increased their presence among the global frontier firms are the United States, Sweden, Finland and the Netherlands. Others which lost ground include Spain and to a lesser extent also France, while the same is true for Italy in manufacturing. While this investigation is preliminary, it shows that using industry-level PPP conversions to get internationally comparable prices and applying the imputations described above can yield some interesting and intuitive cross-country patterns.

5.2 The Impact of Product Market Regulations on Productivity Growth

Following Griffith et al. (2004), Nicoletti and Scarpetta (2003) and Bourlès et al. (2010), but using a wider set of countries (notably, also including non-European ones) and more recent years, we explore the impact of product market regulations on firm-level TFP using a neo-Schumpeterian framework (i.e. catch-up to the frontier) of the following form:

$$\Delta a_{icst} = \alpha_1 \Delta a_{Fst} + \alpha_2 \tilde{a}_{ics,t-1} + \beta RegImpact_{cs,t-1} + \eta_s + \eta_{ct} + \varepsilon_{icst}, \quad (5.1)$$

where Δ is the first difference operator, $a = \ln(TFP)$ is the log of TFP , subscripts i, c, s, t denote firm, country, sector and time, respectively, and the gap to the productivity frontier \tilde{a}_{icst} is defined by

$$\tilde{a}_{icst} = a_{ics} - a_{Fst},$$

where subscript F denotes the frontier firm. The frontier firm, following Arnold et al. (2008), is defined as “the average TFP of the 5% most productive firms in sector s in year t in our sample of countries.” The fixed effects η_s and η_{ct} control for industry-specific and country*time specific (i.e. also including business cycle) effects. The coefficient of the frontier firm’s growth is expected to be positive, $\alpha_1 > 0$, reflecting technological pass-through from the frontier, and the coefficient of the distance from the frontier

⁴⁰More precisely, the top 10% best performing firms are first selected for each industry; then the percentage of firms of a given industry and in a given country is within the top 10% group is calculated; finally, the average across industries is calculated for each country, and those figures are presented in Table 10.

⁴¹The relatively low position of Japan in this ranking probably warrants further examination.

Table 10: The percentage of firms in the global top 10%, by each country, in selected years

		Total market sector																				
		USA	SWE	FIN	FRA	NLD	BEL	DNK	DEU	GBR	AUT	ESP	ITA	KOR	GRC	PRT*	CZE	JPN	POL	SVN	HUN	
2000	21.0%	17.3%	24.8%	12.6%	13.6%	6.4%	9.0%	8.0%	10.1%	9.4%	6.7%	7.2%	3.4%	3.0%	2.6%	2.0%	4.4%					
2005	34.0%	25.4%	21.2%	19.4%	13.8%	12.3%	10.2%	11.5%	10.2%	8.2%	7.6%	5.2%	3.8%	3.7%	2.5%	1.8%	1.6%					
2008	35.5%	27.3%	25.2%	19.4%	16.5%	14.0%	12.0%	11.0%	12.0%	9.2%	5.7%	9.8%	7.7%	4.3%	4.5%	2.5%	2.3%					
N	1,178	10,471	4,903	38,182	2,944	11,359	6,952	28,260	27,547	1,450	50,385	40,277	20,484	5,618	945	10,426	41,369	13,879	2,165	5,462		
		Manufacturing																				
		USA	SWE	FRA	BEL	FIN	NLD	GBR	DEU	AUT**	ITA	DNK	ESP	KOR	GRC	JPN	POL	CZE	SVN	HUN	PRT*	SVK**
2000	35.2%	21.1%	22.3%	22.8%	13.4%	12.0%	10.1%	10.2%	10.1%	5.3%	7.7%	4.8%	4.7%	1.7%	1.3%	2.2%	3.3%	4.3%				
2005	54.5%	32.5%	24.3%	21.8%	20.7%	15.3%	14.5%	11.7%	7.4%	6.4%	5.9%	3.7%	2.7%	2.7%	1.9%	1.9%	1.4%	1.2%				
2008	53.4%	36.1%	18.9%	26.4%	31.3%	20.5%	19.8%	11.6%	10.1%	8.0%	7.1%	3.9%	6.7%	2.5%	4.0%	3.0%	2.6%	2.1%				
N	683	2,880	10,939	3,288	1,718	776	8,043	10,953	533	20,058	2,042	14,917	12,719	2,209	13,034	5,065	4,491	1,044	2,183	386	1,102	

* Portugal has missing price indices in 2008.

** Austria and Slovakia have too few observations in the early 2000s to be included.

Note: the percentages are calculated as follows: Percentage in Top 10%_{ct} = $\frac{\text{number of firms with } TFP_{ct} > TFP_t^{p90}}{\text{total number of firms}_{ct}}$ where TFP_t^{p90} denotes the 90th percentile of TFP value across all countries, for each 2-digit NACE industry. A higher value than 10 means the country is overrepresented in the global top 10% of firms. Countries are ranked according to their 2005 values, from the best to the least well represented one in the global top 10%.

N shows the number of observations in 2005 for which the (externally imputed) Solow-residual based TFPs can be calculated.

The TFP measure used here is the Solow residual with cross-country average labour shares per sector, calculated using external imputations for each firm and country in order to maximize cross-country coverage. Currency conversion is based on industry-PPPs (see Section B). Resampling weights are not available for Japan and Korea, and in order to include them, resampling is not applied here. Instead, firms with at least 20 employees are included only, where firm coverage is more balanced. The implicit assumption is that the representation of firms in ORBIS with respect to TFP is not systematically different across countries for the set of observations where we can use to compute externally imputed Solow-residual based TFPs.

Table 11: The effect of regulations on firm-level productivity growth

Left hand side variable: Δa_{icst}	TFP measures			
	Superlative index measure		Solow residual measure	
Δa_{Fst}	0.103*** (5.6)	0.103*** (5.6)	0.0999*** (5.29)	0.0998*** (5.29)
$a_{ics,t-1} - a_{Fs,t-1}$	-0.148*** (-12.92)	-0.148*** (-12.96)	-0.155*** (-13.65)	-0.155*** (-13.72)
$RegImpact_{cs,t-1}$	-0.317*** (-3.66)		-0.392*** (-3.88)	
$RegImpact_{PO}_{cs,t-1}$		-0.258*** (-3.31)		-0.324*** (-3.62)
Country*year fixed effects	yes	yes	yes	yes
Industry fixed effects	yes	yes	yes	yes
R^2	0.173	0.171	0.182	0.178
N	25,590,249	25,590,249	25,941,974	25,941,974
Number of countries	18	18	18	18

t -statistics in parentheses

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Note: Δa_{icst} denotes firm-level TFP growth, Δa_{Fst} is growth of the global frontier (defined as the average of the top 5% globally most productive firms of the same industry and year), $a_{ics,t-1} - a_{Fs,t-1}$ is the lagged difference from the global frontier, $RegImpact_{cs,t-1}$ stands for the knock-on impact of regulations in upstream, services industries (see Conway and Nicoletti (2006)), and $RegImpact_{PO}_{cs,t-1}$ is a variant of this measure incorporating the effect of the prevalence Public Ownership. Subscripts c, s, i, t denote country, sector, firm and year, respectively. See equation 5.1 for details. Standard errors are clustered at country*industry cells. Resampling weights are used (see details in Section 2.4). The estimation sample covers all non-frontier firms in the non-farm, non-financial business sector for the years 1999-2009 for the broadest set of countries where TFP measures using only external imputations can be calculated, resampling weights are available, industry PPP-s can be calculated, and the RegImpact variable is also available (18 countries): AUT, BEL, CZE, DEU, DKN, ESP, FIN, FRA, GBR, GRC, HUN, ITA, NLD, POL, PRT, SWE, SVK, USA. Both TFP measures use uniform, cross-country average labour shares (Solow) or reference values (Superlative Index) in order to ensure international comparability of productivity levels.

is expected to be negative, $\alpha_2 < 0$, reflecting the fact that as a firm gets closer to the frontier, the speed of catching-up slows down. The main variable of interest is $RegImpact_{cs,t-1}$ which is a country * industry * time-varying measure of the knock-on impact of regulations in upstream, services industries.⁴² If its coefficient is significant and negative, it means that regulations hamper productivity growth at the firm-level, even after controlling for the potential catch-up behaviour of firms to the global frontier.

Equation 5.1 is estimated with all non-frontier firms in the sample and using two types of TFP measures best suited to international comparisons: the Superlative index measure with uniform cross-country average reference values for each firm and the Solow residual, with the same factor shares for each country, calculated as the average of factor shares across countries for each industry. In order to maximize cross-country and firm coverage, externally imputed TFP measures are used.

Results are presented in Table 11. They show that the frontier growth and the distance from the frontier are significant and have the expected signs and similar magnitudes across specifications. These results are comparable to what was found by Arnold et al. (2008) using the Amadeus data set for 10 European countries and for the early 2000s. Furthermore, the impact of regulations is also significant and negative, meaning that firm-level TFP growth is slower if regulations are more stringent.

6 Conclusions

Recent OECD research has utilized harmonized cross-country firm level data from ORBIS to explore the contribution of public policies to cross-country differences in productivity, innovation and resource allocation. This paper describes the steps taken before the analysis in order to improve the reliability and international comparability of these firm level data as well as the specific trade-offs involved when using ORBIS for productivity analysis. Moreover, it carries out a wide range of quality and robustness checks of the measures obtained, also highlighting various potential uses of the data. As such, it aims to serve as a reference for current and future firm-level analyses utilizing ORBIS either within or outside the OECD.

Since certain variables are not available or not with wide enough coverage for some countries, most importantly the United States, productivity measures using internal (using other variables from ORBIS) or external imputations (using external data) are required. When the properties of these imputed measures are compared with those of using non-imputed data, for a set of countries where data coverage is relatively high (France, Germany, Great Britain, Italy and Spain), a number of key results emerge:

- The correlations between imputed and non-imputed TFP estimates are reasonably high, both in terms of levels and growth rates, and this relationship is stable over time and across sectors.
- There are some differences across productivity measures in the quality of the imputations: the Solow residual and the estimation based approaches have correlations around 0.8-0.9 with the non-imputed values, for levels, and around 0.6-0.85 for growth rates.
- When comparing the dispersion of productivity estimates calculated using the imputed and non-imputed variables, dispersion levels are lower for externally imputed values, but the relative differences across industries are well preserved.

As a contribution to the literature on productivity measurement, this paper also compared productivity using different methodologies (index numbers and estimation-based approaches for TFP and labour

⁴²The *RegImpact* variable is a continuous measure, developed by the OECD using national sources based on legislation, with higher values showing a stronger potentially negative effect of non-manufacturing regulations on entry and competition downstream sectors, depending on the sector's use of the output of the regulated sectors (energy, transport, telecommunication, retail distribution and professional services). For more details, see Conway and Nicoletti (2006).

productivity), for countries with good data coverage. The estimation based approaches deliver TFP estimates very close to each other, whereas the index numbers are somewhat less closely related.

Finally, the database uses industry-level PPP conversions in order to ensure the comparability of estimates across countries with different currencies. Also, differences in the coverage of firms across countries are taken care of by using resampling weights in order to match the observed industry and size class structure of employment in administrative databases.

To illustrate the potential of the data set for cross-country comparisons and policy analysis, some preliminary results are presented on the global productivity frontier of firms, the catch-up to those firms and the effect of regulations on productivity growth. These results are in an explorative phase, but they seem to be plausible and confirm earlier findings, namely that the global productivity frontier is disproportionately dominated by the most developed and innovative countries (e.g. United States and Sweden), and that productivity growth is hampered by product market regulations.

Despite all the efforts, some limitations remain – some of them more serious than others, depending on the purpose of the analysis where the data is used. First, in cases where resampling weights are high, typically, in the early years (up to 2004), and in countries with generally poor coverage (see Table 3), results need to be treated with more caution. Further, the measurement of entry and especially exit is noisy. As the data vendor BvD continuously broadens and refines its sources, these issues may be better addressed in future vintages of ORBIS.

Appendix

A Estimating Capital Stock at the Firm Level

For calculating capital stock at the firm-level, the standard Perpetual Inventory Method (PIM) is used. It defines the level of real capital stock K_{it} in firm i in year t as

$$K_{it} = K_{i,t-1}(1 - \delta_{it}) + I_{it}, \quad (\text{A.1})$$

where real investments I_{it} are calculated as the difference between the current and lagged book value of fixed tangible assets plus depreciation, deflated by country and industry specific investment deflators

$$I_{it} = (K_{it}^{BV} - K_{i,t-1}^{BV} + DEPR_{it}^{BV}) / PI_t, \quad (\text{A.2})$$

and where K^{BV} and $DEPR^{BV}$ denote the book value of fixed tangible assets and depreciation, respectively (TANGIBLE_FIXED_ASSETS and DEPRECIATION in ORBIS) and PI_t is the investment price deflator at the 2-digit industry level (see more on this in Section B). The depreciation rate is defined as $\delta_{it} = DEPR_{it}^{BV} / K_{i,t-1}^{BV}$.⁴³

For the first observed year of the firm ($t = 0$), where $K_{i,t-1}$ in equation A.1 is not defined, the real capital stock is simply the observed net capital stock deflated by the investment price index:

$$K_{i0} = K_{i0}^{BV} / PI_0. \quad (\text{A.3})$$

B Deflation and Construction of Internationally Comparable Price Levels

B.1 Deflation - Price Indices over Time

The variables in ORBIS are given in thousands of euros, converted from local currencies at the market exchange rate, in each year (see Gonnard and Ragoussis (2013)). In order to be able to compare values over time, adjusted for price changes, we need to deflate the nominal variables using price indices. Similarly to most firm-level data sets, ORBIS does not include firm-level price indices.⁴⁴ Hence, following the standard practice, external data sources are used to deflate nominal values: 2-digit industry deflators from the OECD STAN database, separately for output, value added, intermediate inputs and investments.

However, these deflators refer to national currencies, hence before they can be applied to nominal variables in ORBIS (which are in thousands of euros), those nominal variables need to be converted back to the original currency values. For this, the variable called *EXCHANGE_RATE* is used from ORBIS, which is the euro / local currency conversion rate for each country and year, to be denoted by $(\text{€}/\$)_{ct}$.

⁴³Since different asset types (i.e. structures and equipment) are summed together in ORBIS, but their depreciation rates differ, this constructed depreciation rate reflects differences in the asset structure (i.e. firms with more equipment will have higher depreciation rates)

⁴⁴It is a general problem in productivity measurement that in most firm-level data sets, firm-level prices are not observed. Hence what researchers usually are able to measure is revenue based productivity (*TFPR*), and not quantity based productivity (*TFPQ*). The former is also influenced by idiosyncratic firm-level prices, and as such, it captures a combination of market power and productivity (Syverson, 2011).

Thus we obtain for each variable X_{it} , for firm i and year t , a conversion from the euro based value ($X_{it}^{\text{€}}$) to the local currency value ($X_{it}^{\text{\$}}$):

$$X_{it}^{\text{\$}} = \frac{X_{it}^{\text{€}}}{(\text{€}/\text{\$})_{ct}}. \quad (\text{B.1})$$

Then at the second stage, deflation over time is implemented:

$$X_{it}^{\text{\$,}t_0} = \frac{X_{it}^{\text{\$}}}{P_{cjt}^{t_0}} \quad (\text{B.2})$$

where $P_{cjt}^{t_0}$ refers to the appropriate deflator from OECD STAN (e.g. value added, output, etc.) in year t with reference year t_0 in country c in industry j . This procedure will ensure that the *growth rates* of the variables used for productivity calculations (value added, output, investment, etc.) are not distorted by price changes.⁴⁵

B.2 Price-level Corrections across Countries

In order to be able to compare *productivity levels* across countries, as for instance in Arnold et al. (2008), Bartelsman et al. (2008) and Griffith et al. (2009) when defining the global frontier firms, we need to use a conversion which corrects for price-level differences across countries. The ICOP 1997 database from the Groningen Growth and Development Center, by Timmer et al. (2007), was prepared using a comprehensive and well documented approach.

To give a quick background on industry-level PPPs, there are two approaches for their measurement (Pilat, 1996): (1) The output approach, which is theoretically more preferable as it uses producer prices. But these are more difficult to obtain for services, and are more suited for industries producing rather homogeneous goods as they utilize an “average price”. (2) The other approach is the expenditure approach, which is more widely available across the economy, but needs to be adjusted for transport and distributional costs, taxes and subsidies, international trade and intermediate use. Timmer et al. (2007) select between the two approaches for each industry and country, based on the quality of these adjustments, which in turn depend on the degree of international trade and intermediate use. Their final database contains industry level PPP exchange rates for each country and each industry (mostly at the 2-digit level).⁴⁶ The PPPs there refer to 1997 values in the form of local currency / euro conversion rates which we denote here by $(\text{\$/€})_{c,j,PPP97}$. Comparable values across countries, for the reference year of 1997, can thus be obtained by

$$X_{i,1997}^{\text{€},PPP97} = \frac{X_{i,1997}^{\text{\$}}}{(\text{\$/€})_{c,j,PPP97}}.$$

B.3 Combining Deflation over Time and across Countries

In order to get comparable values both across countries and over time, we need to combine the two steps, i.e. deflating over time and correcting for price level differences across countries. Hence we choose 1997 as the reference year in equation B.2, i.e. $t_0 = 1997$, and then apply both the across-time deflator $P_{c,j,t}^{1997}$

⁴⁵Where 2-digit values are not available, higher aggregation levels are used. If those are also not available for certain deflators, value added deflators are used, as they are the most widely available.

⁴⁶From the set of countries included in the TFP calculations, Norway and Switzerland do not have values in the ICOP PPP database, hence they are left out of the calculations using industry PPP-s. They are still included in another set of calculations, using a simpler and more crude way of correcting cross-country price differences, see Section B.3 below.

and the conversion rates $(\$/\epsilon)_{c,j,PPP97}$.⁴⁷

$$X_{it}^{\epsilon,PPP97} = \frac{X_{it}^{\$/P_{cjt}^{1997}}}{(\$/\epsilon)_{c,j,PPP97}} = \frac{X_{it}^{\$,1997}}{(\$/\epsilon)_{c,j,PPP97}}. \quad (\text{B.3})$$

An alternative, less sophisticated way than using industry level PPPs is to convert the local currency nominal values using observed market exchange rates to a common currency. It is more acceptable to use in industries prone to strong international competition (typically, manufacturing). In order to mitigate the impact of fluctuating exchange rates on the productivity numbers, we fix local currency vs. euro exchange rates at the average of 2005, corresponding to the middle of the sample period 1999-2009.⁴⁸

C Balancing the SDBS Database

The Structural and Demographic Business Statistics (SDBS) of the OECD is built up using administrative data from national sources on the number of firms, employees and other economic variables, by country, year, industry and firm size class cells. It is used as the benchmark when constructing the resampling weights by cells. For some cells, however, values are missing. Below are the main steps of balancing, i.e. filling up missing values:

- There are two employment variables available in SDBS: number of employees and number of persons engaged. The difference can be quite substantial for small firm size classes, as the latter captures also managers. However, the number of employees is closer to the definition in ORBIS, hence more preferable for our uses, but is not available in some cases where the number of persons engaged is. Therefore, we predict the number of employees for those country*firms*sizeclass*year cells where it is not available, by using country, sizeclass*industry and year fixed effects and interacting these fixed effects with the number of persons engaged.
- Next, the industry aggregation is chosen uniformly across countries, in such a way that it yields the most widely available data (somewhat less detailed than 2-digit, see Table E.4 in the Appendix, the column on NACE 2-digit industry groups).
- Then, for missing years, we use the average values for the period later than 2005 or earlier than 2006, depending on to which interval the actual missing year belongs to. For instance, values for 2009 will get the average available years starting from 2005, possibly those of 2005-2008 but perhaps only those of 2005-2006. Similarly, values for 2001 will be the averages before and including 2006. This approximation is expected to work well, since most of the variation in SDBS is along the cross-sectional and not in the time dimension. In other words, it is more important to get a good sense of the average share of industries and firm-size classes than to exactly follow their evolution over time.

⁴⁷During the actual implementation, we first construct a series of deflators P_{cjt}^{PPP97} :

$$P_{cjt}^{PPP97} = (\$/\epsilon)_{c,j,PPP97} \times P_{cjt}^{1997},$$

and apply those to the nominal values (in local currencies) in ORBIS:

$$X_{it}^{\epsilon,PPP97} = X_{it}^{\$/P_{cjt}^{PPP97}}.$$

This procedure is equivalent to what is written in equation B.3.

⁴⁸A relative advantage of this approach is that it is available for those countries (Switzerland and Norway) that are not in the ICOP-PPP database. In the final TFP database, both approaches are implemented, see the database description in Section F of the Appendix.

- Finally, for completely missing sizeclass categories in certain countries and industries, we apply the average share from other countries and industries.⁴⁹ In principal, we aim to get the most detailed size-class categories. These are, for most countries, the following intervals, defined by the number of employees: 1-9, 10-19, 20-49, 50-249, 250+. One exception is the United States, where the sizeclass definitions are 1-9, 10-19, 20-99, 100-499, 500+⁵⁰

D Comparing Production Function Parameters

Regarding the estimation based approaches, one expects the semi-parametric methods to improve on the potential bias of the input coefficients. As expected, compared to OLS, the coefficient of employment is lower when we control for endogeneity of inputs (i.e. when using Wooldridge (2009)). Another interesting result is that the sum of the coefficients is usually less than unity for the semi-parametric approaches, implying decreasing returns to scale. Note that this holds both for the coefficients based on non-imputed data and for the ones based on imputations.⁵¹

⁴⁹For instance, if in a 2-digit industry, the information referring to the smallest size-class is not available, than the average share of small firms in that industry is assumed, using all other countries for the same industry.

⁵⁰Japan and Korea are problematic countries because some industries are completely missing or present only with leaving out the smallest firm category. See for details the SDBS database on the OECD website. Therefore, resampling weights there should be used with caution.

⁵¹These results are available on request. Coefficients for further inspection are stored next to the database in separate files, see Section F of the Appendix on the structure of the database.

E Additional Tables

Table E.1: Comparing the availability of different productivity measures (2008, market sector)

Country Code	Country	Labor Prod. using Turnover	TFP, non imputed	Non-imputed TFP / Labor productivity	TFP, internally imputed	Internally imputed TFP / Labor productivity	TFP, externally imputed*	Externally imputed TFP / Labor productivity*
ESP	Spain	503,334	464,600	92.30%	466,197	92.62%	508,424	101.01%
PRT	Portugal	231,378	204,603	88.43%	199,859	86.38%	239,208	103.38%
ITA	Italy	345,577	300,704	87.02%	322,347	93.28%	334,927	96.92%
SWE	Sweden	166,810	112,808	67.63%	133,532	80.05%	150,550	90.25%
GRB	Great Britain	45,832	27,465	59.93%	45,492	99.26%	48,603	106.05%
FIN	Finland	37,971	19,885	52.37%	31,918	84.06%	33,445	88.08%
SVN	Slovenia	9,724	4,531	46.60%	9,296	95.60%	9,463	97.32%
NZL	New Zealand	18	8	44.44%	10	55.56%	10	55.56%
KOR	Korea	80,712	32,889	40.75%	59,398	73.59%	63,167	78.26%
NOR	Norway	3,366	1,180	35.06%	1,667	49.52%	2,099	62.36%
CZE	Czech Republic	62,536	21,414	34.24%	35,411	56.62%	38,636	61.78%
FRA	France	398,776	134,300	33.68%	301,115	75.51%	334,499	83.88%
DEU	Germany	105,487	35,255	33.42%	35,107	33.28%	35,811	33.95%
JPN	Japan	220,388	70,988	32.21%	80,731	36.63%	117,404	53.27%
HUN	Hungary	14,072	4,276	30.39%	14,019	99.62%	14,419	102.47%
POL	Poland	48,449	14,044	28.99%	29,566	61.02%	32,285	66.64%
BEL	Belgium	34,852	9,708	27.85%	116,520	334.33%	118,521	340.07%
AUT	Austria	9,969	1,750	17.55%	1,681	16.86%	1,807	18.13%
SVK	Slovak Republic	40,711	5,948	14.61%	7,357	18.07%	7,805	19.17%
LUX	Luxembourg	467	56	11.99%	61	13.06%	64	13.70%
AUS	Australia	1,067	99	9.28%	150	14.06%	895	83.88%
IRL	Ireland	665	27	4.06%	28	4.21%	34	5.11%
CHE	Switzerland	8,024	315	3.93%	369	4.60%	387	4.82%
TUR	Turkey	3,577	69	1.93%	72	2.01%	84	2.35%
NLD	Netherlands	3,853	62	1.61%	4,944	128.32%	5,292	137.35%
ISR	Israel	402	4	1.00%	4	1.00%	393	97.76%
ISL	Iceland	1,347	3	0.22%	936	69.49%	2,213	164.29%
GRC	Greece	19,544	4	0.02%	4	0.02%	17,324	88.64%
DNK	Denmark	11,647	2	0.02%	39,106	335.76%	42,892	368.27%
EST	Estonia	30,077	2	0.01%	19,728	65.59%	21,661	72.02%
USA	United States	5,036,569	1	0.00%	1	0.00%	2,214	0.04%
CAN	Canada	640,965	-	0.00%	-	0.00%	12	0.00%
CHN	China	340	-	0.00%	-	0.00%	1	0.29%
MEX	Mexico	43,539	-	0.00%	-	0.00%	1	0.00%
BRA	Brazil	7,297	14	0.19%	14	0.19%	40	0.55%
CHN	China	182,206	108	0.06%	110	0.06%	1,645	0.90%
IND	India	259	-	0.00%	-	0.00%	1	0.39%
RUS	Russia	620,827	-	0.00%	-	0.00%	-	0.00%
ZAF	South Africa	142	38	26.76%	38	26.76%	42	29.58%
Total (OECD)		8,162,045	1,467,000	17.97%	1,956,626	23.97%	2,184,550	26.76%
Mean (OECD)		240,060	43,147	26.52%	57,548	64.12%	64,251	79.33%
Total (All)		8,972,776	1,467,160	16.35%	1,956,788	21.81%	2,186,278	24.37%
Mean (All)		230,071	37,619	23.81%	50,174	56.59%	56,058	69.96%

Note: OECD countries are ranked by the 5th column (the ratio of the sample for which the simplest TFP measures without imputations can be calculated to the Labour Productivity sample).

The column referring the TFP shows the number of observations where TFP can be calculated, using only the following variables (ORBIS variables in parentheses): value added (ADDED_VALUE), capital (TANGIBLE_FIXED_ASSETS and DEPRECIATION), employment (EMPLOYEES). Internally imputed TFP means that the value added numbers are calculated as the sum of labour costs (COSTS_EMPLOYEES) and capital costs (more precisely, earnings before interest, taxes, depreciation and amortization, EBITDA). Externally imputed TFP means that the labour costs are imputed externally as the product of firm-specific employment and country, industry, year specific average labour costs using OECD STAN. The data underlying the table refers to the non-farm business sector (NACE Rev 1.1 codes 15-74) for the years 1999-2009, and is derived from filtering out accounts referring to less than complete calendar years and consolidated accounts, using flags prepared by the Statistics Directorate of the OECD (STD) (Gonnard and Ragoussis, 2013). Note that in the June 2011 version of OECD-ORBIS, data in the United States and Canada contain only single-establishment firms. The actual numbers in the TFP database may be lower than here if price indices are not available or data problems prevent the calculation of TFP measures for some firms.

Table E.2: Comparing the availability of different productivity measures (2002, market sector)

Country Code	Country	Labor Prod. using Turnover	TFP, non imputed	Non-imputed TFP / Labor productivity	TFP, internally imputed	Internally imputed TFP / Labor productivity	TFP, externally imputed*	Externally imputed TFP / Labor productivity*
PRT	Portugal	1,379	1,307	94.78%	1,325	96.08%	1,366	99.06%
ITA	Italy	238,158	193,966	81.44%	200,756	84.30%	227,529	95.54%
FIN	Finland	37,416	30,266	80.89%	34,103	91.15%	35,936	96.04%
ESP	Spain	395,132	315,485	79.84%	351,440	88.94%	373,323	94.48%
SWE	Sweden	126,632	98,442	77.74%	111,199	87.81%	119,189	94.12%
DNK	Denmark	885	617	69.72%	1,686	190.51%	1,761	198.98%
FRA	France	267,910	174,463	65.12%	251,787	93.98%	265,998	99.29%
FBR	Great Britain	57,764	35,405	61.29%	58,153	100.67%	61,650	106.73%
POL	Poland	15,056	8,181	54.34%	10,874	72.22%	10,999	73.05%
KOR	Korea	67,010	29,267	43.68%	51,214	76.43%	53,695	80.13%
SVK	Slovak Republic	2,320	1,008	43.45%	2,053	88.49%	2,187	94.27%
CZE	Czech Republic	21,242	6,583	30.99%	16,122	75.90%	18,064	85.04%
JPN	Japan	93,200	28,080	30.13%	30,310	32.52%	65,792	70.59%
NOR	Norway	66,526	17,388	26.14%	19,389	29.14%	57,027	85.72%
IRL	Ireland	114	28	24.56%	24	21.05%	47	41.23%
BEL	Belgium	42,332	8,057	19.03%	94,179	222.48%	95,354	225.25%
AUS	Australia	692	110	15.90%	114	16.47%	191	27.60%
SVN	Slovenia	6,199	905	14.60%	3,559	57.41%	3,619	58.38%
CHE	Switzerland	2,354	297	12.62%	327	13.89%	352	14.95%
LUX	Luxembourg	200	21	10.50%	20	10.00%	21	10.50%
NZL	New Zealand	14	1	7.14%	2	14.29%	12	85.71%
DEU	Germany	64,986	4,464	6.87%	4,358	6.71%	4,533	6.98%
HUN	Hungary	6,598	146	2.21%	2,544	38.56%	2,634	39.92%
ISL	Iceland	576	12	2.08%	448	77.78%	889	154.34%
ISR	Israel	52	1	1.92%	1	1.92%	40	76.92%
NLD	Netherlands	7,727	94	1.22%	7,997	103.49%	8,429	109.09%
AUT	Austria	9,139	79	0.86%	78	0.85%	81	0.89%
TUR	Turkey	872	4	0.46%	4	0.46%	22	2.52%
EST	Estonia	17,075	19	0.11%	11,715	68.61%	12,735	74.58%
USA	United States	19,606	1	0.01%	2	0.01%	3,115	15.89%
CAN	Canada	9,829	-	0.00%	-	0.00%	83	0.84%
CHL	Chile	21	-	0.00%	-	0.00%	8	38.10%
GRC	Greece	18,515	-	0.00%	-	0.00%	16,339	88.25%
MEX	Mexico	143	-	0.00%	-	0.00%	67	46.85%
BRA	Brazil	111	25	22.52%	26	23.42%	64	57.66%
CHN	China	149,185	60	0.04%	61	0.04%	385	0.26%
IND	India	310	2	0.65%	2	0.65%	2	0.65%
RUS	Russia	83	-	0.00%	-	0.00%	-	0.00%
ZAF	South Africa	203	36	17.73%	34	16.75%	40	19.70%
Total (OECD)		1,597,674	954,697	59.76%	1,265,783	79.23%	1,443,087	90.32%
Mean (OECD)		46,990	28,079	28.22%	37,229	54.77%	42,444	73.29%
Total (All)		1,747,566	954,820	54.64%	1,265,906	72.44%	1,443,578	82.61%
Mean (All)		44,809	24,483	25.66%	32,459	48.79%	37,015	65.90%

Note: Same as below Table E.1

Table E.3: Correlation between levels and growth rates of value added: imputed vs. non-imputed values

ESP	$\ln(\text{Value added})$			$\Delta \ln(\text{Value added})$		
	2002	2005	2008	2002	2005	2008
internal	0.988	0.984	0.990	0.886	0.865	0.945
external	0.928	0.927	0.922	0.626	0.639	0.713
external_refined	0.960	0.960	0.960	0.672	0.698	0.752
ITA	$\ln(\text{Value added})$			$\Delta \ln(\text{Value added})$		
	2002	2005	2008	2002	2005	2008
internal	0.979	0.989	0.987	0.837	0.814	0.856
external	0.949	0.954	0.913	0.721	0.589	0.657
external_refined	0.973	0.978	0.953	0.794	0.656	0.711
GBR	$\ln(\text{Value added})$			$\Delta \ln(\text{Value added})$		
	2002	2005	2008	2002	2005	2008
internal	0.986	0.982	0.982	0.806	0.800	0.813
external	0.940	0.941	0.937	0.631	0.613	0.625
external_refined	0.969	0.970	0.966	0.676	0.679	0.661
FRA	$\ln(\text{Value added})$			$\Delta \ln(\text{Value added})$		
	2002	2005	2008	2002	2005	2008
internal	0.985	0.983	0.980	0.762	0.741	0.724
external	0.946	0.940	0.930	0.557	0.530	0.499
external_refined	0.970	0.966	0.958	0.616	0.579	0.523
DEU	$\ln(\text{Value added})$			$\Delta \ln(\text{Value added})$		
	2002	2005	2008	2002	2005	2008
internal	0.968	0.974	0.971	0.724	0.736	0.776
external	0.922	0.932	0.925	0.412	0.506	0.567
external_refined	0.952	0.952	0.959	0.411	0.609	0.642
Country average*	$\ln(\text{Value added})$			$\Delta \ln(\text{Value added})$		
	2002	2005	2008	2002	2005	2008
internal	0.980	0.981	0.975	0.778	0.806	0.794
external	0.925	0.922	0.914	0.578	0.598	0.595
external_refined	0.959	0.956	0.950	0.645	0.680	0.638

* Unweighted average across all those countries where the non-imputed value added measure is available and the number of observations is at least 200 for the specific year. Internal, external and external refined refer to the type of imputations which are used to construct imputed value added measures (see equations 2.2-2.4 in Section 2.3). The numbers stay almost identical when the value added measures are regressed first on industry*year fixed effects and the residual variations are correlated.

Table E.4: Industry Breakdowns by their NACE Codes and Descriptions

<u>2-digit industries (nace2)</u>	<u>2-digit industry groups (nace2_groups)</u>	<u>1-letter industry-groups (nace1)</u>
15 Food and beverages	15_16 FOOD, BEVERAGES AND TOBACCO	D MANUFACTURING
16 Tobacco		
17 Textiles	17_19 TEXTILES, TEXTILE, LEATHER AND FOOTWEAR	
18 Wearing Apparel, Dressing And Dying Of Fur		
19 Leather, leather and footwear		
20 WOOD AND OF WOOD AND CORK	20 WOOD AND OF WOOD AND CORK	
21 Pulp, paper and paper	21_22 PULP, PAPER, PAPER, PRINTING AND PUBLISHING	
22 Printing, publishing and reproduction		
23 Coke, refined petroleum and nuclear fuel	23_25 CHEMICAL, RUBBER, PLASTICS AND FUEL	
24 Chemicals and chemical		
25 Rubber and plastics		
26 OTHER NON-METALLIC MINERAL	26 OTHER NON-METALLIC MINERAL	
27 Basic metals	27_28 BASIC METALS AND FABRICATED METAL	
28 Fabricated metal		
29 MACHINERY, NEC	29 MACHINERY, NEC	
30 Office, accounting and computing machinery	30_33 ELECTRICAL AND OPTICAL EQUIPMENT	
31 Electrical machinery and apparatus, nec		
32 Radio, television and communication equipment		
33 Medical, precision and optical instruments		
34 Motor vehicles, trailers and semi-trailers	34_35 TRANSPORT EQUIPMENT	
35 Other transport equipment		
36 Manufacturing nec	36_37 MANUFACTURING NEC; RECYCLING	
37 Recycling		
40 ELECTRICITY AND GAS	40_41 ELECTRICITY, GAS AND WATER SUPPLY	E ELECTRICITY, GAS AND WATER SUPPLY
41 WATER SUPPLY		
45 CONSTRUCTION	45 CONSTRUCTION	F CONSTRUCTION
50 Sale, maintenance and repair of motor vehicles and motorcycles; retail sale of fuel	50_52 WHOLESALE AND RETAIL TRADE	G WHOLESALE AND RETAIL TRADE
51 Wholesale trade and commission trade, except of motor vehicles and motorcycles		
52 Retail trade, except of motor vehicles and motorcycles; repair of household goods		
55 HOTELS AND RESTAURANTS	55 HOTELS AND RESTAURANTS	H HOTELS AND RESTAURANTS
60 Other Inland transport	60_63 TRANSPORT AND STORAGE	I TRANSPORT AND STORAGE AND COMMUNICATION
61 Other Water transport		
62 Other Air transport		
63 Other Supporting and auxiliary transport activities; activities of travel agencies		
64 POST AND TELECOMMUNICATIONS	64 POST AND TELECOMMUNICATIONS	
65 Financial intermediation, except insurance and pension funding	65_67 FINANCIAL INTERMEDIATION	J FINANCIAL INTERMEDIATION
66 Insurance and pension funding, except compulsory social security		
67 Activities related to financial intermediation		
70 Real estate activities	70_74 REAL ESTATE, RENTING AND BUSINESS ACTIVITIES	K REAL ESTATE, RENTING AND BUSINESS ACTIVITIES
71 Renting of machinery and equipment		
72 Computer and related activities		
73 Research and development		
74 Other business activities		

Note: The classification is based on NACE Rev 1.1. The middle column is based on the grouping used in the EU-KLEMS database O'Mahony and Timmer (2009). Since 2008, NACE Rev 2 has been introduced, and firms who started appearing in ORBIS after that were converted using conversion tables from NACE Rev 2 to NACE Rev 1.1, available from Eurostat at http://epp.eurostat.ec.europa.eu/portal/page/portal/nace_rev2/correspondence_tables. Firms appearing in the database before 2008 have their original NACE Rev 1.1 values stored and used throughout the paper and the resulting productivity database.

Table E.5: The average of resampling weights by broad sectors (2005)*

	D	E	F	G-H	I	K	Average
AUT	38.8	8.9	177.5	81.5	38.8	115.3	79.0
BEL	1.4	1.0	1.3	1.5	1.1	2.4	1.6
CZE	3.4	1.6	4.8	4.3	2.4	5.1	4.2
DEU	30.3	5.1	105.7	54.1	27.2	56.3	46.6
DNK	3.3	18.5	3.9	2.8	1.8	3.7	3.3
EST	1.9	2.1	1.8	1.4	1.2	1.6	1.6
ESP	1.7	1.1	2.9	1.9	1.1	2.4	2.1
FIN	1.5	1.6	1.7	1.2	1.1	1.5	1.4
FRA	2.4	1.8	2.6	1.5	1.4	2.6	2.0
GBR	11.7	2.0	32.5	20.2	5.8	19.7	18.2
FRC	4.5	11.7	28.0	7.4	4.5	7.6	7.7
HUN	21.2	6.2	51.0	39.4	7.6	127.5	39.7
ITA	5.7	1.6	13.4	10.9	3.0	15.6	9.4
JPN	15.7						15.7
KOR	11.9			13.8	8.0	7.8	11.7
NLD	27.4	5.0	42.2	35.3	11.7	60.3	40.8
NOR	50.7	110.5	56.2	41.8	30.0	35.7	41.7
POL	14.9	3.1	31.1	23.2	9.3	33.3	20.9
PRT	118.4	12.0	231.6	132.9	35.2	207.3	147.5
SWE	1.3	1.3	1.3	1.1	1.1	1.2	1.2
SVN	2.4	3.2	5.7	2.3	1.2	7.7	3.1
SVK	4.0	1.7	5.4	3.9	1.3	7.0	4.4
USA	683.0	86.8	5485.2	14123.5	5140.7	5340.1	5192.4
Average	10.9	4.8	7.9	26.3	29.6	22.6	26.4

* The sectoral description can be found in Table E.4.

Note: Entries show the average of the constructed resampling weights for 2005, for the Solow-residual (using external imputations). For Korea and Japan, some industries and size classes are not in the SDDBS database which are used for constructing the weights.

Table E.6: Rank Correlations across TFP measures based on Imputed and Non-Imputed variables
Average over ESP, ITA, GBR, FRA, DEU

<i>TFP measures</i>	Rank correlations for $\ln(TFP)$				<i>TFP measures</i>	Rank correlations for TFP growth			
	Solow	Sup.In.	OLS	Wooldridge		Solow	Sup.In.	OLS	Wooldridge
<i>imputation types</i>					<i>imputation types</i>				
internal	0.97	0.92	0.95	0.94	internal	0.91	0.89	0.90	0.83
external	0.83	0.66	0.69	0.75	external	0.69	0.64	0.64	0.56
external, refined	0.89	0.74	0.81	0.80	external, refined	0.71	0.66	0.65	0.59

Note: rank correlations are based for observations in the year 2008 for the non-farm business sector. TFP growth is defined as the first differences of $\ln(TFP)$. For Great Britain, non-imputed Wooldridge type estimates are not available due to the lack of the availability of proxies for intermediate inputs. For a description of the imputation types, see Section 2.3.

F Structure of the Resulting ORBIS-TFP Database

The data in the TFP-database uses the non-farm business sector (NACE Rev 1.1 codes 15-74) for the years 1993-2010, and is derived from filtering out accounts referring to less than complete calendar years and consolidated accounts, using flags prepared by the Statistics Directorate of the OECD (STD) (Gonnard and Ragoussis, 2013). Most countries and observations are in the period 1999-2009, the year 2010 is not complete in the June 2011 version of ORBIS. The number of observations for each country, year and imputation type, for the Superlative Index measure, are contained in the *TFP_Sample.xlsx* file in the root directory of the TFP database.

F.1 The File Structure

The database containing the productivity estimates and resampling weights is structured in the following way:

1. "core data", containing the most important productivity measures, using such an industry breakdown which is most appropriate (Directory: *Core\Data*)
 - Index-number approaches (in parentheses the level of industry breakdown where the average factor shares are calculated)
 - Solow-residuals with cross-country average labour shares (computed at the 2-digit industry level)
 - two types of Superlative-index measures, using country-level means and cross-country means as reference values, (computed at the 2-digit industry level) ⁵²
 - Estimation based approaches (in parentheses the level of industry breakdown where homogeneity of the production function parameters are assumed):
 - OLS (2-digit industry groups)
 - Wooldridge (2-digit industry groups)
2. "additional data", containing additional productivity measures (Directory: *Additional\Data*)
 - Index-number approaches (in parentheses the level of industry breakdown where the average factor shares are calculated):
 - Solow-residuals with country-level labour shares (computed at the 2-digit industry level)
 - Superlative-index measures, using country-level means (at the 2-digit industry level) as reference firms
 - Estimation based approaches (in parentheses the level of industry breakdown where homogeneity of the production function parameters are assumed):
 - OLS (2-digit level)
 - OLS with firm fixed effects (2-digit level, 2-digit industry groups)
 - Levinsohn Petrin (2003) (2-digit level, 2-digit industry groups)
 - Wooldridge (2-digit level)
3. labour productivity data, containing only turnover-based labour productivity measures, for the widest available set of firms.

⁵²The measure using cross-country means as reference firms are stored only in the large, cross-country data files.

As benchmark, labour productivity measures (revenue based and value added based) are also included in the core and additional databases.

The index number approaches assume homogeneity at the 2-digit level, but for estimation based approaches, it is more advisable to use a higher level of aggregation in order to have more observations when running the estimations country-by-country and industry-by-industry. Therefore, the classification in EU-KLEMS (Timmer et al., 2007) is used, where the non-farm business sector is classified into industry-groups by merging together some 2-digit industries (which typically have few firms) (see Table E.4 for details). The additional files contain estimates for the 2 digit level, but that comes at a cost of some countries with fewer firms in the database will have many industries missing due to a low number of observations to run the estimations.

Each of the estimations come with 4 types of imputations: (1) no imputation (2) internal imputation (i.e. using the alternative definition of value added as the sum of costs of capital and labour), (3) external imputations (4) external imputations using wage predictions (see Section 2.3).

All of the estimates are stored in country-by-country files. For the core estimates, large cross-country databases can also be found, containing all productivity numbers as in the country-by-country files, and in addition, also the Superlative index measures which use the cross-country, firm-level averages, for each industry, as the reference firm.

Furthermore, each estimations are constructed by using all firms with at least one employees and for the subsample of firms with at least 20 employees. The latter subsample has the advantage that the potentially weaker data quality of small firms does not influence the results, and also, the coverage of very small firms is usually weaker in most of the countries in ORBIS. Also, currencies are either converted to a common currency using industry-level PPP values from Timmer et al., 2007 or by using the 2005 average exchange rates between local currencies and the euro. For analysing developments in industries which are prone to strong international competition (typically manufacturing), the latter approach may be more preferable, as the first approach uses industry PPP numbers which refer to 1997.⁵³

Reflecting all of these possibilities, the structure of the file names is the following:

*<2-letter country code>_MinSize_<minimum of mean employment>_ ...
<conversion type>_<date of creating the file>.dta,*

- where the set of countries included in the database is marked in Table 5. Also, large cross-country files are created in the core database, with their first characters being *cross_co*. Only these files contain Superlative Index measures with world averages, and they also contain other measures which are included in the country-by-country files.
- *minimum of mean employment* can take values:
 - 0: no restriction, except the essential one that at least 1 employees are needed for each firm
 - 20: the minimum of mean employment, for each firm, over its observed lifetime, should be at least 20
- *conversion type* can take values:
 - *euro_2005*: for using 2005 average market exchange rates between euro and local currencies

⁵³There is no more up to date data available, to the best of our knowledge. I am grateful for Colin Webb for providing information on this.

- *PPP97_chained*: for using the ICOP industry-level PPP values (Timmer et al., 2007) in 1997 and using local currency price indices over time to get to actual values for the sample period 1999-2009

- *date of creating the file* is *01May2012*.

Normally, one would prefer to use the files with no employment restrictions and which uses the PPP conversions. However, if the question at hand may be sensitive to cross-country level comparisons, as a robustness check, the *euro_2005* versions can also be used. Also, if one wants to focus only on firms with probably better quality of data (more than 20 employees), then that sample can be used (which is also much smaller hence estimations run much faster).

Besides the Stata data files with the productivity estimates, the database contains also small Stata files containing the coefficient values, standard errors and number of observations used for the estimation, for each industry where the estimation is run. They are located in the directories *Core\Coefficients* or *Additional\Coefficients* and have the following file format:

```
<2-letter country code>_<conversion type>_<minimum of mean employment>_<estimation type>_ ...
<industry breakdown level>_<imputation type>_<date of creating the file>.dta,
```

- *estimation type* can take values of *OLS*, *W* (stands for Wooldridge) (in the core database), *OLS_FE*, *LevPet* (only in the additional database)
- *industry breakdown level* can take values of *nace2_groups* (estimations run for each 2-digi industry group, shown in Table E.4) and *nace2* (estimations run for each 2-digit industries - carries the risk of unreliable / impossible estimation for certain industries in countries with low number of observations)
- *imputation type* can take values of 1 (no imputations used), 2 (only internal imputations used), 3 (only external imputations are used), 4 (only refined external imputations are used). For a description on imputations, see Table 2 and Section 2.3 in the main text.

and the other dimensions are the same as for the file names containing the estimates.

Finally, detailed estimation outputs are collected in the log files with the following formats:

```
TFP_estim_<2-letter country code>_<date of creating the file>.txt
```

```
TFP_stats_<2-letter country code>_<date of creating the file>.txt,
```

where the former contains detailed estimation outputs for each round of estimations, and the latter contains descriptive statistics which compares different productivity measures along several dimensions. The tables in this document (Tables 6, 9) are prepared using these outputs.

F.2 Variables in the Productivity Data Files

The content of the data files (the .dta files under *Core\Data* and *Additional\Data*) is summarized in Table F.1. Note that the estimation based measures are residuals of the estimated production functions, and estimations are run only if the number of observations is at least 50 for the given country and industry. In other cases, the estimation is not carried out and the productivity variable is empty for the particular country and industry. Furthermore, if either labour or capital coefficients are non-positive, the productivity numbers are missing. The number of such cases is negligible for countries with a reasonable

number of observations (i.e. at least around 10,000). Minimizing the number of cases in as many countries as possible is the main motivation for recommending 2-digit industry group breakdowns for the estimation based approaches and including those in the core data set.

Table F.1: Variables in the ORBIS Productivity Database

	Variable name	Description
Identifiers	<i>COUNTRY_CODE</i>	2-letter country code
	<i>OECD_ID</i>	Unique firm id constructed by STD
	<i>YEAR</i>	Year
	<i>nace2</i>	2-digit NACE Rev 1.1 industry code
	<i>nace2_groups</i>	2-digit NACE Rev 1.1 industry groups, based on EU-KLEMS groupings
	<i>sector</i>	1-letter NACE Rev 1.1
Economic variables for the productivity calculations	<i>OPERATING_REV_TURNOVER</i>	Revenues or turnover
	<i>EMPLOYEES</i>	Number of employees
	<i>orig_ADDED_VALUE</i>	Original value added (as given in OECD-ORBIS)
	<i>ADDED_VALUE_2</i>	Added value, imputed using the definition as the sum of labour (<i>orig_COSTS_EMPLOYEES</i>) and capital income (EBITDA) (internal imputation)
	<i>ADDED_VALUE_3</i>	Added value, imputed using the definition as for <i>ADDED_VALUE_2</i> but using <i>COSTS_EMPLOYEES_3</i> (average wages from OECD STAN at the 2-digit level; external imputation)
	<i>ADDED_VALUE_4</i>	Added value, imputed using external average wages at the 2-digit level and predictors of wage-differences
	<i>orig_COSTS_EMPLOYEES</i>	Total labour costs (as given in OECD-ORBIS)
	<i>COSTS_EMPLOYEES_3</i>	Total labour costs, imputed using external average wages at the 2-digit level
	<i>COSTS_EMPLOYEES_4</i>	Total labour costs, imputed using external average wages at the 2-digit level and predictors of wage-differences
	<i>ICAPITAL</i>	Log of real tangible capital (constructed using the PIM-method)
	<i>VALP</i>	value added deflator (at the 2-digit NACE Rev 1.1 level) from OECD STAN
Productivity variables*	<i>LP_VA_1</i>	Labour productivity, using value added and no imputations
	<i>LP_TR</i>	Labour productivity, using turnover
	<i>LP_VA_2</i>	Labour productivity, using value added with internal imputations
	<i>LP_VA_3</i>	Labour productivity, using value added with external imputations
	<i>LP_VA_4</i>	Labour productivity, using value added with external imputations and wage-differential predictions
Only in the Core database	<i>TFP_nace2_Solow_CO_1</i>	Total factor productivity (TFP), obtained as a Solow-residual, using 2-digit NACE Rev 1.1, country level labour cost shares from OECD STAN, with no imputations
	<i>TFP_nace2_Solow_W_1</i>	TFP, obtained as a Solow-residual, using 2-digit NACE Rev 1.1, cross-country average labour cost shares from OECD STAN, with no imputations
	<i>TFP_nace2_SupIn_CO_1</i>	TFP, Superlative-index measure, using 2-digit NACE Rev 1.1, country level averages as reference values, with no imputations
	<i>TFP_nace2_groups_OLS_1</i>	TFP, OLS, using 2-digit industry groups, no imputations
	<i>TFP_nace2_groups_W_1</i>	TFP, Wooldridge (2009), using 2-digit industry groups, no imputations
	<i>TFP_nace2_Solow_CO_2</i>	The same as above but with internal imputations
	<i>TFP_nace2_Solow_CO_3</i>	The same as above but with external imputations
	<i>TFP_nace2_Solow_CO_4</i>	The same as above but with refined external imputations
Only in the Core, cross country database	<i>TFP_nace2_SupIn_W_1_0</i>	TFP, Superlative-index measure, using 2-digit NACE Rev 1.1, cross-country average levels as reference values, with no imputations
	<i>TFP_nace2_SupIn_W_2_0</i>	The same as above but with internal imputations
	<i>TFP_nace2_SupIn_W_3_0</i>	The same as above but with external imputations
	<i>TFP_nace2_SupIn_W_4_0</i>	The same as above but with refined external imputations
	<i>TFP_nace2_SupIn_W_1_1</i>	TFP, Superlative-index measure, using 2-digit NACE Rev 1.1, cross-country average levels as reference values, with no imputations, resampling weights applied already at the stage of constructing the Superlative Index
	...	The same as above, for all types of imputations
Only in the Additional database	<i>TFP_nace2_groups_Solow_CO_1</i>	Total factor productivity (TFP), obtained as a Solow-residual, using 2-digit NACE Rev 1.1, country level labour cost shares from OECD STAN, with no imputations
	<i>TFP_nace2_groups_Solow_W_1</i>	TFP, obtained as a Solow-residual, using 2-digit industry groups, cross-country average labour cost shares from OECD STAN, with no imputations
	<i>TFP_nace2_groups_SupIn_CO_1</i>	TFP, Superlative-index measure, using 2-digit industry groups, country level averages as reference values, with no imputations
	<i>TFP_nace2_OLS_1</i>	TFP, OLS, using 2-digit industries, no imputations
	<i>TFP_nace2_W_1</i>	TFP, Wooldridge (2009), using 2-digit industries, no imputations
	<i>TFP_nace2_OLS_FE_1</i>	TFP, OLS with firm fixed effects, using 2-digit industries, no imputations
	<i>TFP_nace2_groups_OLS_FE_1</i>	TFP, OLS with firm fixed effects, using 2-digit industry groups, no imputations
	<i>TFP_nace2_LevPet_1</i>	TFP, Levinsohn-Petrin (2003), with firm fixed effects, using 2-digit industries, no imputations
	<i>TFP_nace2_groups_LevPet_1</i>	TFP, Levinsohn-Petrin (2003), with firm fixed effects, using 2-digit industry groups, no imputations
	<i>TFP_nace2_groups_Solow_CO_2</i>	The same as above but with internal imputations
	<i>TFP_nace2_groups_Solow_CO_3</i>	The same as above but with external imputations
<i>TFP_nace2_groups_Solow_CO_4</i>	The same as above but with refined external imputations	

* All productivity variables have a corresponding variable with the same name but with a "w_" prefix, which contain resampling weights appropriate for the specific productivity variable. Each productivity measure is calculated and stored in logs.

F.3 Auxiliary Data

In a separate directory (*Auxiliary*), further external information is stored which are used during the calculations:

- Conversion tables between NACE Rev 2 to NACE Rev 1.1
- Price indices and average labour costs (country*industry*year) from OECD STAN (see Section B)
- The industry-level ICOP PPP database (See Timmer et al., 2007)
- Employment (country*industry*sizeclass*year) from SDBS, and its filled-up version (see Section 2.4)
- *RegImpact* variable (country*industry*year), provided by the Economics Department of the OECD

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